NONPARAMETRIC STOCHASTIC BANDITS

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BMGE, MIT, Intelligent Data Analysis, Nov 19, 2013

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• Y_{kt} – payoff of arm k when chosen the *t*th time, $1 \le k \le K$

- For k fixed, Y_{kt} is an i.i.d. sequence
- $\mu_k = \mathbb{E}[Y_{kt}]$
- $\mu^* = \max_k \mu_k$
- For $k \neq k'$, Y_{kt} and $Y_{k't'}$ are independent
- $J_{ ext{bad}} = \{k \, | \mu_k < \mu^*\}$, set of "bad" arms
- $J_{\text{good}} = \{k | \mu_k = \mu^*\}$, set of "good" arms
- I_t^A choice of arm at time t by some allocation rule A
- $T_{kt}^{\mathcal{A}} = \sum_{s=1}^{t} \mathbb{I}_{\{l_s^{\mathcal{A}} = k\}}$ (# of choosing *k*)

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• $T_{I_t,t} = #$ of pulls of arm I_t till time t

- Y_{lt,Tlut} =payoff in the t-step
- Payoff/-Loss in n steps is

$$L_n(\mathcal{A}) = \sum_{t=1}^n Y_{l_t, T_{l_t, t}}$$

Expected regret in *n* steps:

$$R_n(\mathcal{A}) \stackrel{\text{def}}{=} \sup_{\mathcal{A}'} \mathbb{E}\left[L_n(\mathcal{A}')\right] - \mathbb{E}\left[L_n(\mathcal{A})\right].$$

- Goal: Minimize regret!
- Constraint: Distributions of the payoffs are unknown.

This is stochastic bandit.

There is non-stochastic: $\{Y_{kt}\}_{t\geq 1}$ is not i.i.d. random, but any individual sequence, \mathbb{E} [] is replaced by sup over them. Variation of special case of *prediction with expert advice*.

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• Exercise #1: Expected payoff

$$\mathbb{E}\left[L_n(\mathcal{A})\right] = \sum_{k=1}^{K} \mu_k \mathbb{E}\left[T_{kn}\right] \le n\mu^*$$

Hint: Use Wald's identity. *T_{kn}* is stopping time w.r.t. ...
Exercise #2:

 $\sup_{\mathcal{A}'} \mathbb{E}\left[R_n(\mathcal{A}')\right] = n\mu^*$

• Let $\Delta_k = \mu^* - \mu_k$. Hence:

$$R_n(\mathcal{A}) = n\mu^* - \sum_{k=1}^K \mu_k \mathbb{E}[T_{kn}] = \sum_{k \in J_{\text{bad}}} \Delta_k \mathbb{E}[T_{kn}].$$

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$$R_n(\mathcal{A}) = n\mu^* - \sum_{k=1}^K \mu_k \mathbb{E}\left[T_{kn}\right] = \sum_{k \in J_{\text{bad}}} \Delta_k \mathbb{E}\left[T_{kn}\right].$$

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• Exercise #1: Expected payoff

$$\mathbb{E}\left[L_n(\mathcal{A})\right] = \sum_{k=1}^{K} \mu_k \mathbb{E}\left[T_{kn}\right] \le n\mu^*$$

Hint: Use Wald's identity. *T_{kn}* is stopping time w.r.t. ...
Exercise #2:

 $\sup_{\mathcal{A}'} \mathbb{E}\left[R_n(\mathcal{A}') \right] = n \mu^*$

• Let $\Delta_k = \mu^* - \mu_k$. Hence:

$$R_n(\mathcal{A}) = n\mu^* - \sum_{k=1}^{K} \mu_k \mathbb{E}[T_{kn}] = \sum_{k \in J_{\text{bad}}} \Delta_k \mathbb{E}[T_{kn}].$$

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WALD'S IDENTITIES

A r.v. *T* is a stopping time w.r.t. a sequence $\{Y_t\}$ of r.v.'s, if for each *t*, $\mathbb{I}_{\{T \le t\}}$ depends only on Y_1, \ldots, Y_t .

LEMMA (WALD'S IDENTITIES — SPECIAL CASE)

Let $\{Y_t\}$ be an i.i.d. sequence of r.v.'s, T be a stopping time w.r.t. $\{Y_t\}$, and $\mathbb{E}[T] < \infty$. If $\mathbb{E}[|Y_1|] < \infty$ then

$$\mathbb{E}\left[\sum_{t=1}^{T} Y_t\right] = \mathbb{E}\left[Y_1\right] \mathbb{E}\left[T\right].$$

If $\mathbb{E}\left[\, Y_{1}^{2} \, \right] < \infty$ then

$$\mathbb{E}\left[\left(\sum_{t=1}^{T} Y_t - T\mathbb{E}\left[Y_1\right]\right)^2\right] = \operatorname{Var}\left[Y_1\right]\mathbb{E}\left[T\right].$$

WALD'S IDENTITIES — GENERAL

T is a stopping time w.r.t. a filtration $\{\mathcal{F}_t\}$, if for each *t*, $\{T \leq t\} \in \mathcal{F}_t$.

LEMMA (WALD'S IDENTITIES)

Let { \mathcal{F}_t } be a filtration and { Y_t } be \mathcal{F}_t -adapted i.i.d. sequence of r.v.'s. Assume that \mathcal{F}_t and σ ({ $Y_s : s \ge t + 1$ }) are independent, *T* is a stopping time w.r.t. \mathcal{F}_t , and $\mathbb{E}[T] < \infty$. If $\mathbb{E}[|Y_1|] < \infty$ then

$$\mathbb{E}\left[\sum_{t=1}^{T} Y_t\right] = \mathbb{E}\left[Y_1\right] \mathbb{E}\left[T\right].$$

If $\mathbb{E}\left[\, Y_1^2 \, \right] < \infty$ then

$$\mathbb{E}\left[\left(\sum_{t=1}^{T} Y_t - T\mathbb{E}\left[Y_1\right]\right)^2\right] = \operatorname{Var}\left[Y_1\right]\mathbb{E}\left[T\right].$$

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- 2 REGRET
- **3** ϵ -GREEDY POLICIES
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Notation:

$$\overline{\mathbf{Y}}_{kt} = \frac{1}{t} \sum_{t'=1}^{t} \mathbf{Y}_{kt'}$$

• Assumption: $0 \le Y_{kt} \le 1$ (in the rest of the talk!)

-GREEDY

- Initialization: Choose all arms 1, ..., Konce.
- At time 1 choose arm with the maximal payoff with probability 1 — g otherwise an arm uniformly

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• ϵ -greedy choice:

$$\mathbb{P}\left(I_t = \operatorname{argmax}_k \overline{Y}_{kt} \middle| \{\overline{Y}_{kt}\}_{1 \le k \le K}\right) = 1 - \epsilon_t.$$

- *ϵ_t* = 0: always choose maximum. Why is this bad?

 Exercise #3: Give a lower bound on the regret for Bernoulli bandits
- $\epsilon_t = 1$ clearly not good
- Fix *ϵ_t* = *ϵ*: regret still linear. Exercise #4: Give a lower bound on the regret for 0 < *ϵ* < 1

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LOGARITHMIC REGRET

IDEA!

In order to achieve logarithmic (cumulative) regret, the probability of not selecting the best looking arm in step *t* should be $\approx 1/t$, since $\sum_{t=1}^{n} 1/t \approx \ln n!$

Theorem (instantions regret bound [Auer et al., 2002])

Let $\Delta_{\min} = \min_{j \in J_{bad}} \Delta_j$. Let $\epsilon_t = \min(1, \frac{5K}{\Delta_{\min}^2 t})$ time dependent. If $n \ge 5K/\Delta_{\min}$ then

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THEOREM (INSTANTANEOUS REGRET BOUND [AUER ET AL., 2002])

Let $\Delta_{\min} = \min_{j \in J_{bad}} \Delta_j$. Let $\epsilon_t = \min(1, \frac{5K}{\Delta_{\min}^2 t})$ time dependent. If $n \ge 5K/\Delta_{\min}$ then

$$\mathbb{P}\left(I_n \notin J_{\text{good}}\right) = O\left(\frac{1}{\Delta_{\min}^2 n}\right) \text{ and } R_n(\mathcal{A}_{\epsilon}) \leq O\left(\frac{1}{\Delta_{\min}^2}\right) \ln n.$$

• Two sources of error:

- Randomization (fine, by design!)
- Not picking an optimal arm when we wanted to; assuming single optimal arm with index k^* , with $I_t = \operatorname{argmax}_i \overline{Y}_{it}$:

$$\mathbb{P}\left(I_t \neq k^*\right) = \mathbb{P}\left(\overline{Y}_{I_t,t} > \overline{Y}_{k^*,t}\right) = \dots$$

- We need to compare the probability that one average is larger than another one
- How to do this? Solution: Law of large numbers: Averages are close to their expected values: $\overline{Y}_{k,t} \approx \mu_k < \mu^* \approx \overline{Y}_{k^*,t}$
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Mild assumptions on X! (no parametric forms)

- Markov: $X \ge 0$ then $\mathbb{P}(X \ge \epsilon) \le \mathbb{E}[X]/\epsilon$
- Now, for any $\phi : \mathbb{R}^+ \to \mathbb{R}^+$ strictly increasing,

 $\mathbb{P}\left(X \geq \epsilon\right) = \mathbb{P}\left(\phi(X) \geq \phi(\epsilon)\right) \leq \mathbb{E}\left[\phi(X)\right]/\phi(\epsilon).$

 Chebyshev: Choose φ(ε) = ε²! Let X be such that Var [X] < ∞. Then using for |X − E [X] |:

$$\mathbb{P}\left(|X - \mathbb{E}\left[X\right]| \ge \epsilon\right) \le \frac{\operatorname{Var}\left[X\right]}{\epsilon^2}.$$

How tight is Chebyshev's inequality??

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CHEBYSHEV'S INEQUALITY FOR AVERAGE

$$\mathbb{P}\left(|\overline{\mathbf{Y}}_{n} - \mathbb{E}\left[\mathbf{Y}_{1}\right]| \geq \epsilon\right) \leq \frac{\operatorname{Var}\left[\mathbf{Y}_{1}\right]}{n\epsilon^{2}}$$

INTUITION: CENTRAL LIMIT THEOREM

$$\mathbb{P}\left(\overline{Y}_n - \mathbb{E}\left[Y_1\right] \ge \epsilon\right) = \mathbb{P}\left(\frac{\sqrt{n}}{\sigma}\left(\overline{Y}_n - \mathbb{E}\left[Y_1\right]\right) \ge \frac{\sqrt{n}}{\sigma}\epsilon\right)$$
$$\to 1 - \Phi\left(\frac{\sqrt{n}}{\sigma}\epsilon\right) \approx e^{-n\epsilon^2/(2\sigma^2)}\frac{\sigma}{\sqrt{n}\epsilon} \approx e^{-n\epsilon^2/(2\sigma^2)}.$$

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exponential \Rightarrow much sharper could be!

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SHARPENING THE BOUNDS

• For $\phi \ge 0$ strictly increasing:

 $\mathbb{P}\left(\boldsymbol{X} \geq \epsilon\right) = \mathbb{P}\left(\phi(\boldsymbol{X}) \geq \phi(\epsilon)\right) \leq \mathbb{E}\left[\phi(\boldsymbol{X})\right] / \phi(\epsilon).$

• Higher moments: $\phi(\epsilon) = \epsilon^q$, $q \ge 2$:

 $\mathbb{P}\left(\left|X - \mathbb{E}\left[X\right]\right| \ge \epsilon\right) \le \mathbb{E}\left[\left|X - \mathbb{E}\left[X\right]\right|^{q}\right] / \epsilon^{q}.$

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Sharpening the bounds/2

• Chernoff's method: $\phi(x) = e^{sx}$, s > 0;

$$\mathbb{P}\left(X \geq t\right) \leq \mathbb{E}\left[e^{sX}\right]e^{-st}$$

and optimize for s!

• Apply to $n\overline{Y}_n = S_n = \sum_{t=1}^n Y_t$:

$$\mathbb{P}(S_n - \mathbb{E}[S_n] \ge t) \le e^{-st} \mathbb{E}\left[e^{s(S_n - n\mathbb{E}[Y_1])}\right]$$
$$= e^{-st} \prod_{t=1}^n \mathbb{E}\left[e^{s(Y_t - \mathbb{E}[Y_1])}\right]$$

• Hoeffding: $\mathbb{E}[X] = 0$, $a \le X \le b$ then $\mathbb{E}[e^{sX}] \le e^{s^2(b-a)^2/8}$

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$$\mathbb{P}\left(X\geq t\right)\leq\mathbb{E}\left[e^{sX}\right]e^{-st}$$

and optimize for s!

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- 2 Regret
- **3** ϵ -GREEDY POLICIES
- 4 HOEFFDING'S INEQUALITY
- **5** Algorithm UCB1
- 6 ANALYSIS OF THE REGRET OF UCB1
- **O** EXTENSIONS
- 8 BIBLIOGRAPHY

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UCB1

- Initialization: Use all arms once
- Step t > K: Use arm with highest index

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Slightly better than [Auer et al., 2002]: tradeoff in p explicit

 Coefficient ∑<sub>i∈J_{bad} 1/Δ_i is large, if many small Δ_i > 0, i.e., hard to distinguish the best arms.
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THEOREM (UCB1 REGRET)

Let $0 \le Y_{it} \le 1$. Then the regret of UCB1 when used with $c_{t,T} = \sqrt{\frac{p \ln t}{2T}}$ and p > 2 satisfies

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HEURISTIC ANALYSIS

Recall: $R_n(A) = \sum_{i \in J_{\text{bad}}} \Delta_i \mathbb{E}[T_{in}]$, hence we bound $\mathbb{E}[T_{in}]$ for bad *i* arms.

Fact 1

If confidence intervals do not fail and $I_t = i$ then

$$\mu^* - \mu_i = \max_j \mu_j - \mu_i \le 2c_{t,T_{it}},$$

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hence $c_{t,T_{it}} \geq \Delta_i/2$.

PROOF BY FIGURE!

GOAL: ASSUMING $I_t = i$, prove $c_{t,T_{it}} \ge \Delta_i/2!$



(Actually, the conclusion holds even if we only have $\mu_i \ge \overline{Y}_{i,T_{it}} - c_{t,T_{it}}$ and $\mu_j \le \overline{Y}_{j,T_{jt}} + c_{t,T_{jt}}$.)

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HEURISTIC ANALYSIS/2

• By Fact 1, with high prob. if $I_t = i$ then $c_{t,T_{it}} \ge \Delta_i/2$, i.e.,

$$\frac{\Delta_i^2}{4} \le c_{t,T_{it}}^2 \sim \frac{p \ln t}{2T_{it}}, \qquad \text{hence} \qquad T_{it} \le \sim \frac{2p \ln t}{\Delta_i^2}.$$

Thus, using $t \le n$, for a bad arm $\mathbb{E}[T_{in}] \le 2p \ln n/\Delta_i^2$, and

$$R_n = \sum_i \Delta_i \mathbb{E} \left[T_{in} \right] \leq \sim \sum_{i \in \mathcal{J}_{\mathrm{bad}}} \frac{1}{\Delta_i} \cdot O(\ln n).$$

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8 BIBLIOGRAPHY

• UCT = UCB applied to searching in Trees [Kocsis and Szepesvári, 2006];

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- Budgeted learning: some costs instead of time steps
- Best arm identification
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FOR FURTHER READING



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