

AIT-BUDAPEST



AQUINCUM INSTITUTE OF TECHNOLOGY

Creativity in  
Computer Science &  
Engineering

# COMPUTATIONAL BIOLOGY and MEDICINE

## Biomarker discovery I.

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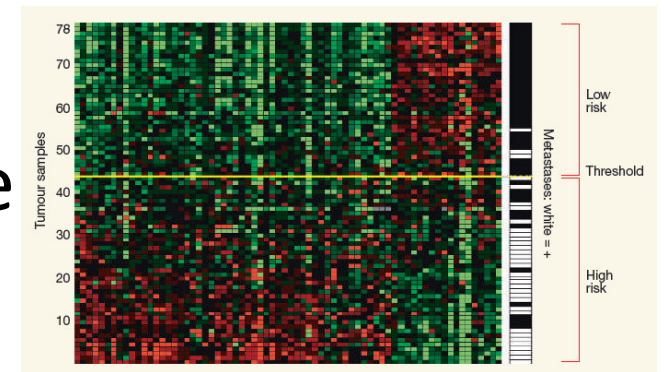
AIT, Budapest 2011. fall

# Overview

- Biomarkers
- Multivariate approach to GAS
- Models: Naïve Bayesian network, logistic regression
- Feature relevance, the feature subset selection problem
- Sufficient and necessary set for diagnosis:
  - Markov blankets
  - Strong relevance
- Identification methods of biomarkers
- The Bayesian statistical approach
- Partial multivariate analysis
- Genagrid
- BayesEye

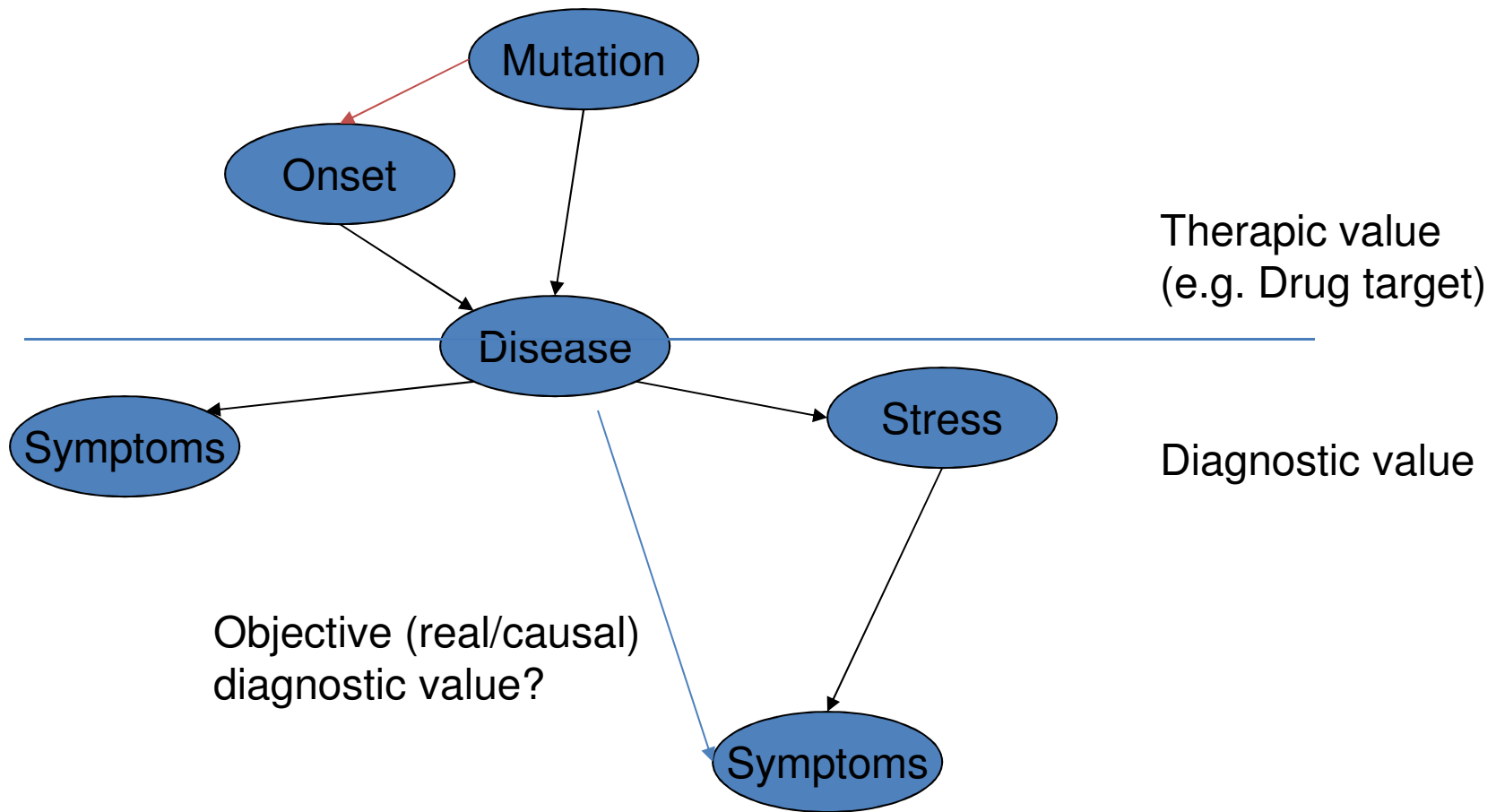
# CA125 and its pretenders

- CA125 is very indicative tumor marker in cancer.
- Missing the mark, 2007, Nature
- MISSING THE MARK: *Why is it so hard to find a test to predict cancer?*, 2011, March
  - Series of proteomic kits... with poor performance.
  - The „prolactin” case

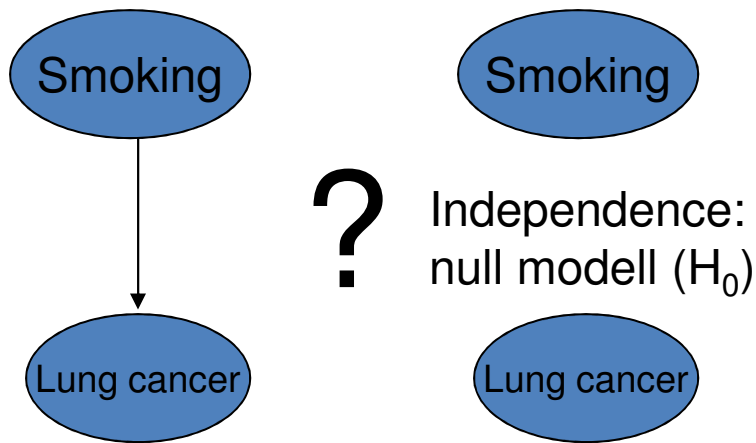


# Causal vs. diagnostic markers

## Direct $\neq$ Causal



# Conditional probabilities, odds, odds ratios



	$\neg S$	S	
$\neg LC$	8	7	15
LC	1	4	5
	9	11	20

Contingency table with marginals

	$\neg S$	S	
$\neg LC$	.4	.35	.75
LC	.05	.2	.25
	.45	.55	

## Conditional probabilities:

$$P(LC | \neg S) = .11 \quad ??? \quad P(LC | S) = .36 \quad ??? \quad P(LC) = .25$$

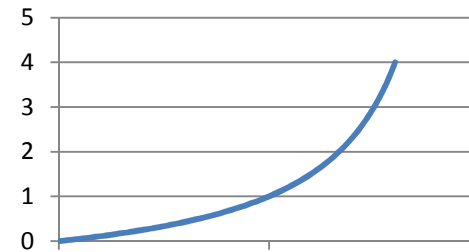
## Odds:

$$[0, 1] \rightarrow [0, \infty]: \text{Odds}(p) = p / (1 - p)$$

$$O(LC | \neg S) = .12 \quad ??? \quad O(LC | S) = .56$$

## Odds Ratio (OR):

$$OR(LC, S) = O(LC | S) / O(LC | \neg S) = 4.6$$



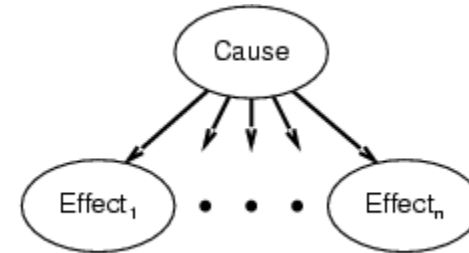
$$\frac{p(x | y)}{p(\neg x | y)} = \frac{p(x | \neg y)^{0.5}}{p(\neg x | \neg y)} = \frac{p(x | y)^1}{p(\neg x | y)} \frac{p(\neg x | \neg y)}{p(x | \neg y)}$$

→ prevalences + odds: joint distribution, e.g.  $P(LC, S) = P(LC | S) P(S)$

# Naive Bayesian network

Assumptions:

1, Two types of nodes: a cause and effects.



2, Effects are conditionally independent of each other given their cause.

## Variables (nodes)

Flu: present/absent

FeverAbove38C: present/absent

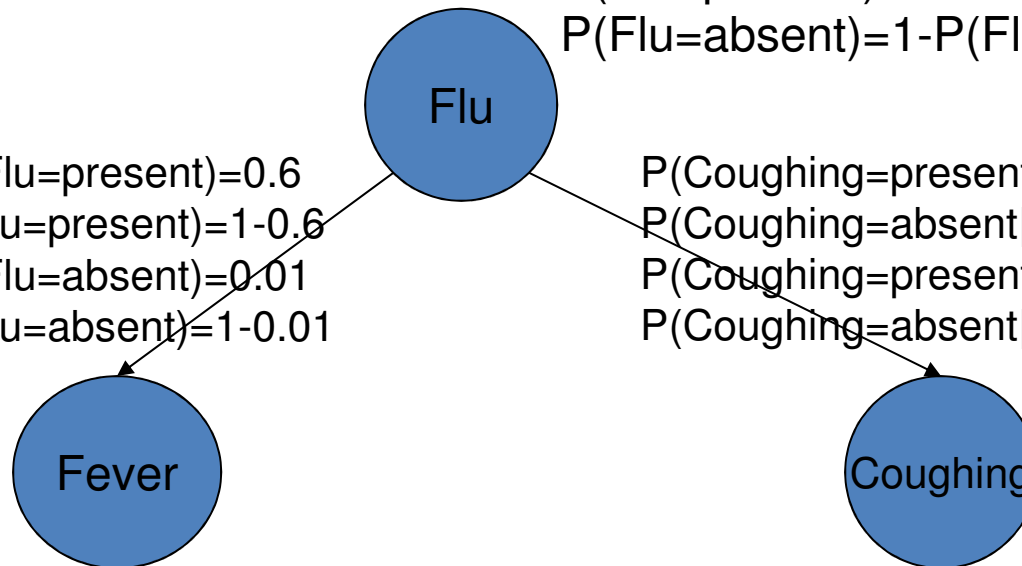
Coughing: present/absent

## Model

$P(\text{Fever}=\text{present}|\text{Flu}=\text{present})=0.6$   
 $P(\text{Fever}=\text{absent}|\text{Flu}=\text{present})=1-0.6$   
 $P(\text{Fever}=\text{present}|\text{Flu}=\text{absent})=0.01$   
 $P(\text{Fever}=\text{absent}|\text{Flu}=\text{absent})=1-0.01$

$P(\text{Flu}=\text{present})=0.001$   
 $P(\text{Flu}=\text{absent})=1-P(\text{Flu}=\text{present})$

$P(\text{Coughing}=\text{present}|\text{Flu}=\text{present})=0.3$   
 $P(\text{Coughing}=\text{absent}|\text{Flu}=\text{present})=1-0.3$   
 $P(\text{Coughing}=\text{present}|\text{Flu}=\text{absent})=0.02$   
 $P(\text{Coughing}=\text{absent}|\text{Flu}=\text{absent})=1-0.02$



# Naive Bayesian network (NBN)

Decomposition of the joint:

$$\begin{aligned} P(Y, X_1, \dots, X_n) &= P(Y) \prod_i P(X_i | Y, X_1, \dots, X_{i-1}) && // \text{by the chain rule} \\ &= P(Y) \prod_i P(X_i | Y) && // \text{by the N-BN assumption} \end{aligned}$$

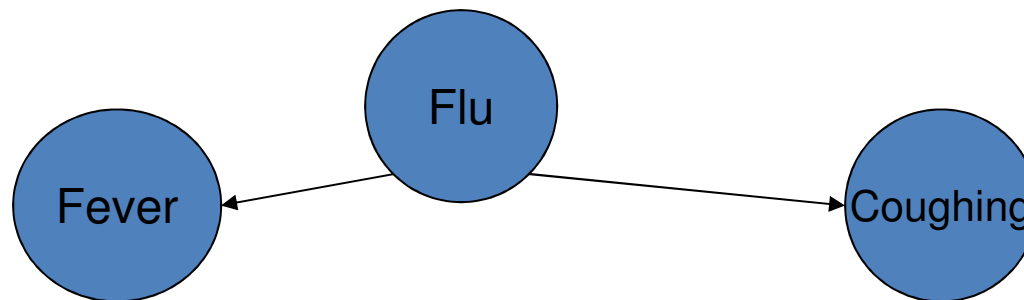
2n+1 parameteres!

Diagnostic inference:

$$P(Y | x_{i1}, \dots, x_{ik}) = P(Y) \prod_j P(x_{ij} | Y) / P(x_{i1}, \dots, x_{ik})$$

If Y is binary, then the odds

$$P(Y=1 | x_{i1}, \dots, x_{ik}) / P(Y=0 | x_{i1}, \dots, x_{ik}) = P(Y=1) / P(Y=0) \prod_j P(x_{ij} | Y=1) / P(x_{ij} | Y=0)$$



$$p(\text{Flu} = \text{present} | \text{Fever} = \text{absent}, \text{Coughing} = \text{present})$$

$$\propto p(\text{Flu} = \text{present}) p(\text{Fever} = \text{absent} | \text{Flu} = \text{present}) p(\text{Coughing} = \text{present} | \text{Flu} = \text{present})$$

# Logistic regression

Assume binary outcomes  $y, \bar{y}$  and predictors  $x_i, \bar{x}_i$ . Logistic regression without interactions can be defined by the odds ratios for the predictors  $x_i, i = 1, \dots, n$  and the bias  $\Psi_0$  ( $x_0 \triangleq 1$ ):

$$\Psi_i = \frac{P(y|x_i)P(\bar{y}|\bar{x}_i)}{P(\bar{y}|x_i)P(y|\bar{x}_i)} \triangleq \exp^{\beta_i}, \Psi_0 = \prod_{i=0}^n \frac{P(y|\bar{x}_i)}{P(\bar{y}|\bar{x}_i)} \triangleq \exp^{\beta_0}.$$

The odds  $P(y|\mathbf{x})/P(\bar{y}|\mathbf{x})$  for a given  $\mathbf{x}$  is defined as

$$P(y|\mathbf{x})/P(\bar{y}|\mathbf{x}) = \prod_{i=0}^n \Psi_i^{x_i} \quad (18)$$

$$\log(P(y|\mathbf{x})/P(\bar{y}|\mathbf{x})) = \sum_{i=0}^n \beta_i x_i \quad (19)$$

$$P(y|\mathbf{x}) = \sigma\left(\sum_{i=0}^n \beta_i x_i\right), \quad (20)$$

where  $\sigma()$  is the logistic sigmoid function  $\sigma(x) = 1/(1 + e^{-x})$ .

$$P(y|\mathbf{x}) = \sigma\left[\sum_{i=0}^n (\beta_i x_i + \sum_{j=1}^n (\beta_{i,j} x_i x_j + \sum_{k=1}^n (\beta_{i,j,k} x_i x_j x_k + \dots)))\right],$$



# Biomarkers and the feature subset selection (FSS) problem

## A probabilistic concept of relevance

**Definition 1.** A feature  $X_i$  is strongly relevant, if there exists some  $x_i, y$  and  $s_i = x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n$  for which  $p(x_i, s_i) > 0$  such that  $p(y|x_i, s_i) \neq p(y|s_i)$ . A feature  $X_i$  is weakly relevant, if it is not strongly relevant, and there exists a subset of features  $S'_i$  of  $S_i$  for which there exists some  $x_i, y$  and  $s'_i$  for which  $p(x_i, s'_i) > 0$  such that  $p(y|x_i, s'_i) \neq p(y|s'_i)$ . A feature is relevant, if it is either weakly or strongly relevant; otherwise it is irrelevant [7, 8].

## A graph-theoretic representation of relevance

**Theorem 1 ([16]).** If distribution  $P$  is stable w.r.t. the DAG  $G$ , then the variables corresponding to the nodes in the boundary of  $Y$ ,  $\text{bd}(Y, G)$  (the parents and children of  $Y$  and other parents of its children) forms a unique and minimal Markov blanket of  $Y$ ,  $\text{MB}_P(Y)$  (the Markov boundary). Furthermore,  $X_i \in \text{MB}_P(Y)$ , if  $X_i$  is strongly relevant.

# Bayesian networks

## Directed acyclic graph (DAG)

- nodes – random variables/domain entities
- edges – direct probabilistic dependencies  
(edges- causal relations)

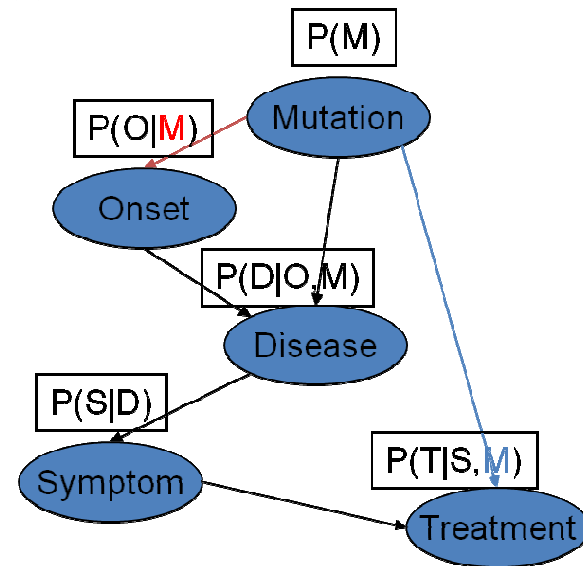
## Local models - $P(X_i | Pa(X_i))$

## Three interpretations:

### 3. Concise representation of joint distributions

$$P(M, O, D, S, T) =$$

$$P(M)P(O|M)P(D|O,M)P(S|D)P(T|S,M)$$



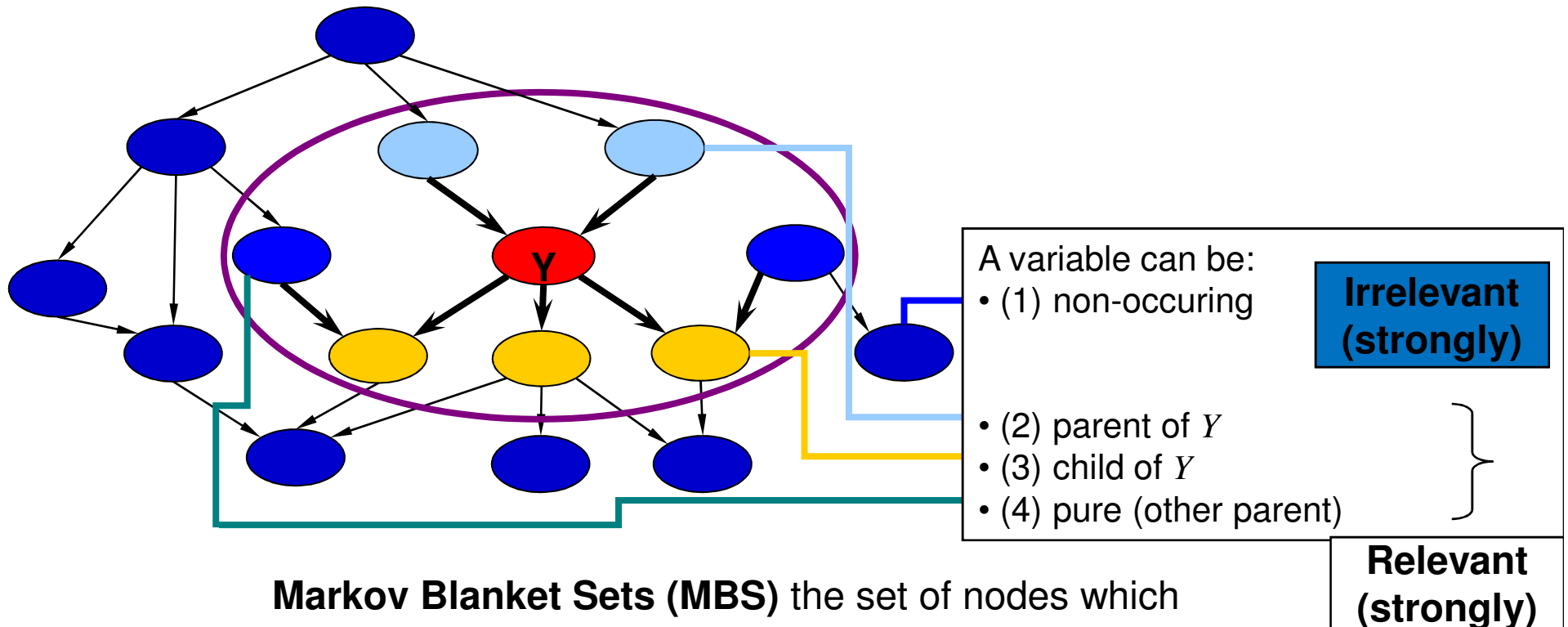
### 1. Causal model

$$M_P = \{I_{P,1}(X_1; Y_1 | Z_1), \dots\}$$

### 2. Graphical representation of (in)dependencies

# The Markov Blanket

A minimal sufficient set for prediction/diagnosis.



**Markov Blanket Sets (MBS)** the set of nodes which probabilistically isolate the target from the rest of the model

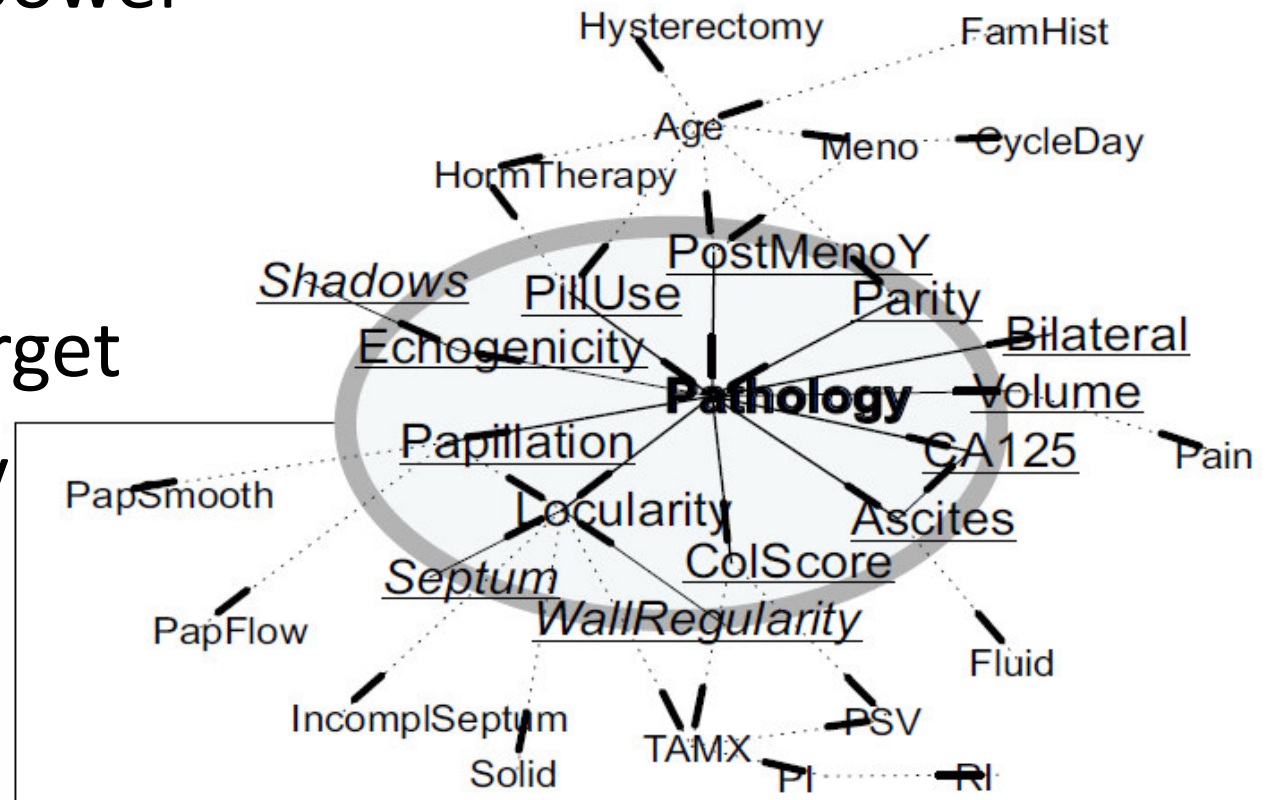
**Markov Blanket Membership (MBM)**

(symmetric) pairwise relationship induced by MBS

# Aspects of biomarkers

„Maximum predictivity, minimum redundancy”

- Predictive power
- Directness
- Causality
- Multiple target
- Uncertainty



# The wrapper approach to FSS

Assume our goal is an „efficient” predictor  $f(\mathbf{X}')=Y$

1. Initialize set  $S$  with a priori good predictors

2. Cycle

- 2.1 Modify  $S$

- E.g. Greedy-univariate: select an additional predictor based on predictive power

- 2.2 Improved predictive power?

- Misclassification rate, AUC, likelihood:  $P(\text{Data} | \text{Model}(S))$

3. Terminate

# The filter approach to FSS („local causal(?) discovery“)

1. Initialize set  $S$  with a priori good predictors
2. Cycle
  1. Expand  $S$ 
    1. E.g. Greedy-univariate: select an additional predictor based on predictive power, which is still relevant to target  $Y$  given  $S$
  2. Decrease  $S$  based on Markov blankets
    1. E.g.  $S-X$  shields  $X$  from target  $Y$
3. Terminate

# The hypothesis testing framework

- Terminology:

- False/true x positive/negative
- Null hypothesis: independence

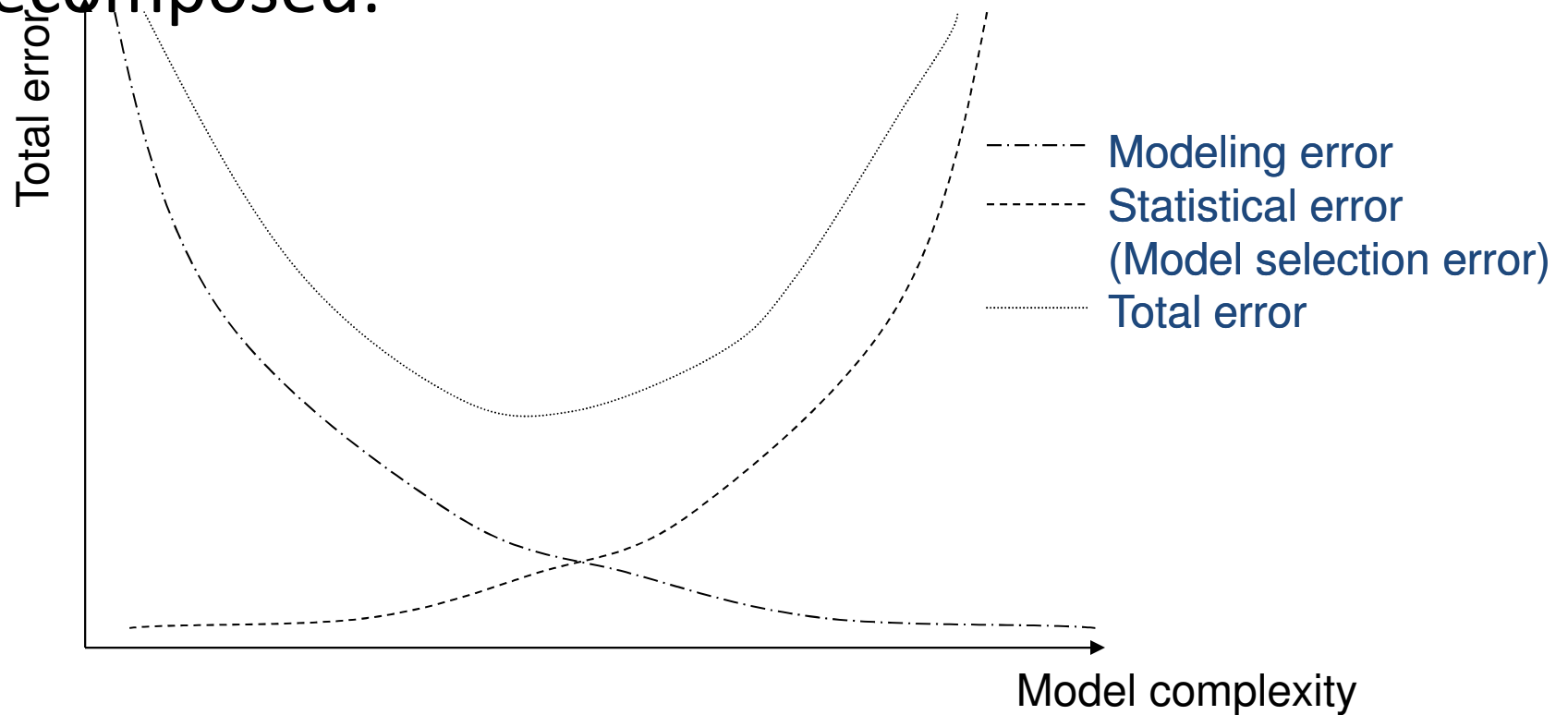
reported	Ref.:0/N	Ref.1/P
0/N	TN	FN
1/P	FP	TP

- Type I error/error of the first kind/ $\alpha$  error/FP:  $p(\neg H_0 | \underline{H}_0)$ 
  - Specificity:  $p(H_0 | \underline{H}_0) = 1 - \alpha$
  - Significance:  $\alpha$
  - p-value: „probability of more extreme observations in repeated experiments”
- Type II error/error of the second kind/ $\beta$  error/FN:  $p(H_0 | \neg \underline{H}_0)$  :
  - Power or sensitivity:  $p(\neg H_0 | \neg \underline{H}_0) = 1 - \beta$

reported	Ref. $\underline{H}_0$	Ref.: $\neg \underline{H}_0$
$H_0$		Type II
$\neg H_0$	Type I („false rejection”)	

# The bias-variance dilemma

- For a given sample size the error is decomposed:





# Bayes rule, Bayesianism

„all models are wrong, but some are useful”

$$p(X | Y) = \frac{p(Y | X) p(X)}{p(Y)}$$

A scientific research paradigm

$$p(\textit{Model} | \textit{Data}) \propto p(\textit{Data} | \textit{Model}) p(\textit{Model})$$

A practical method for inverting causal knowledge to diagnostic tool.

$$p(\textit{Cause} | \textit{Effect}) \propto p(\textit{Effect} | \textit{Cause}) \times p(\textit{Cause})$$

# Bayesian prediction

In the frequentist approach: Model identification (selection) is necessary

$$p(\textit{prediction} \mid \textit{data}) = p(\textit{prediction} \mid \textit{BestModel}(\textit{data}))$$

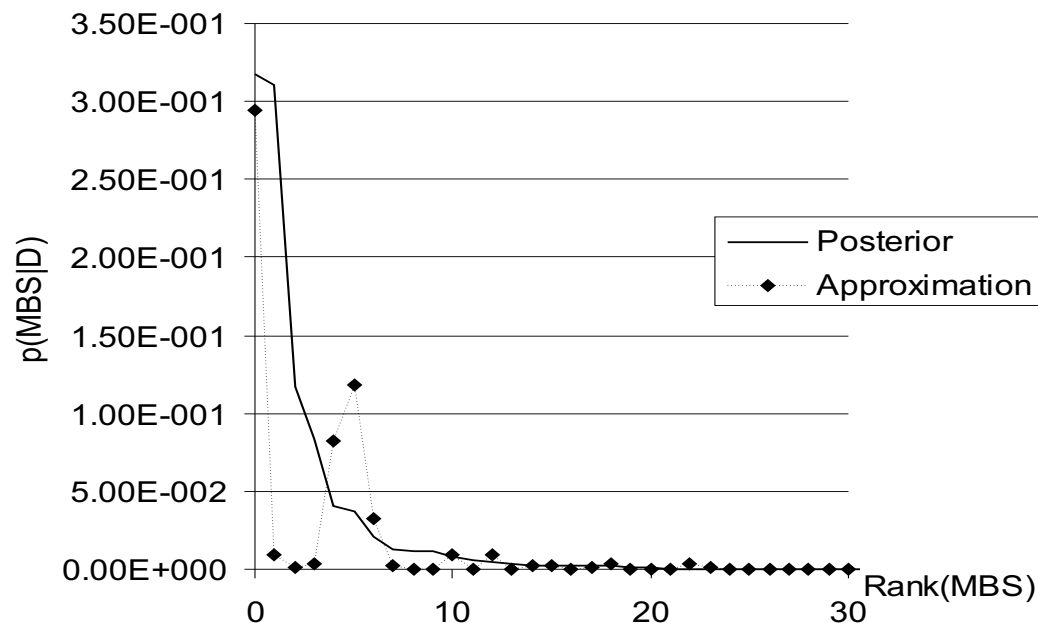
In the Bayesian approach models are weighted

$$p(\textit{prediction} \mid \textit{data}) = \sum_i p(\textit{pred.} \mid \textit{Model}_i) p(\textit{Model}_i \mid \textit{data})$$

Note: in the Bayesian approach there is no need for model selection

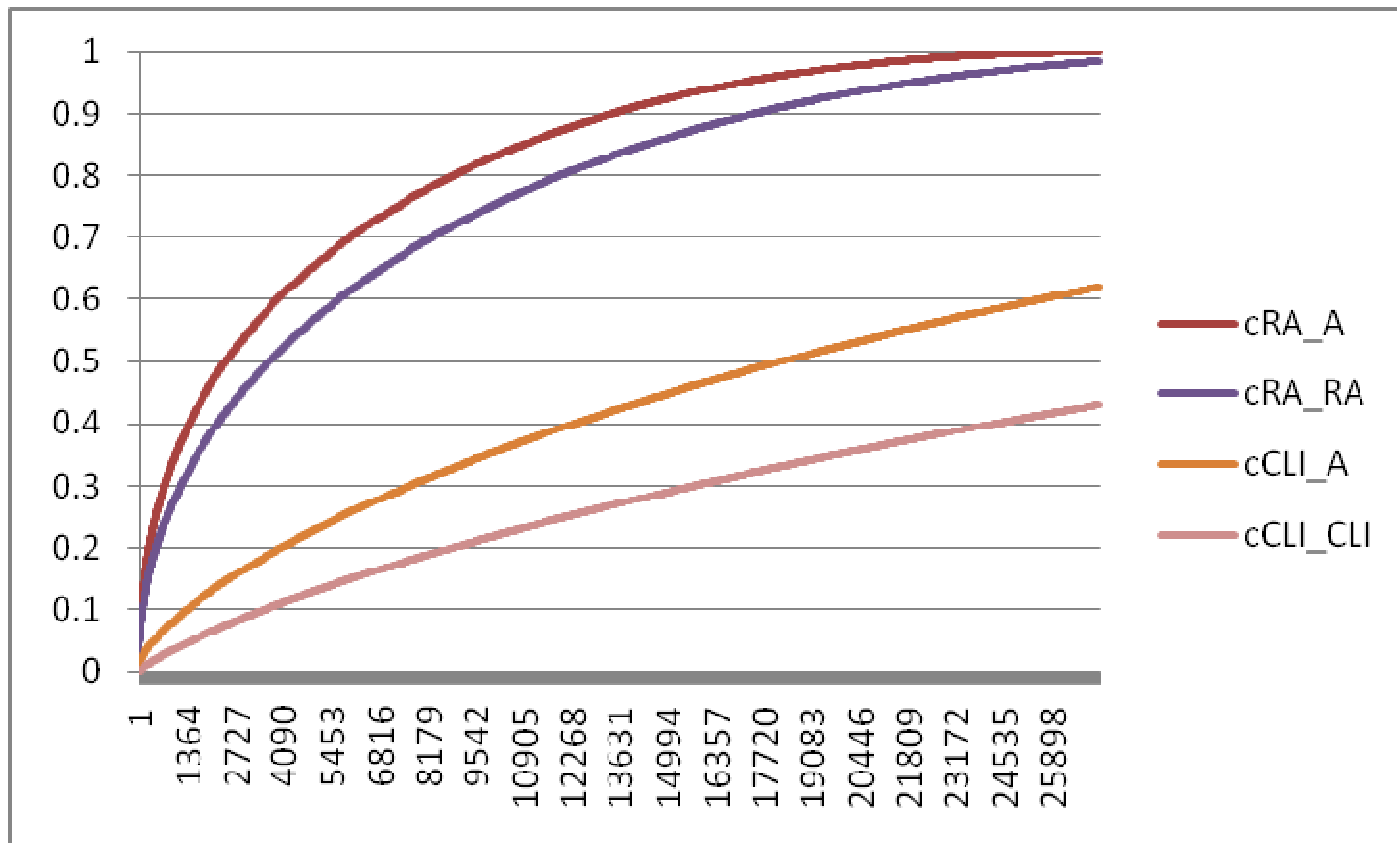
# Posterior for complete sets

- High-level of uncertainty in multivariate analysis
- There are stable sub-parts (e.g., subset, subgraphs)
- Results for target variables and for certain SNPs could be aggregated

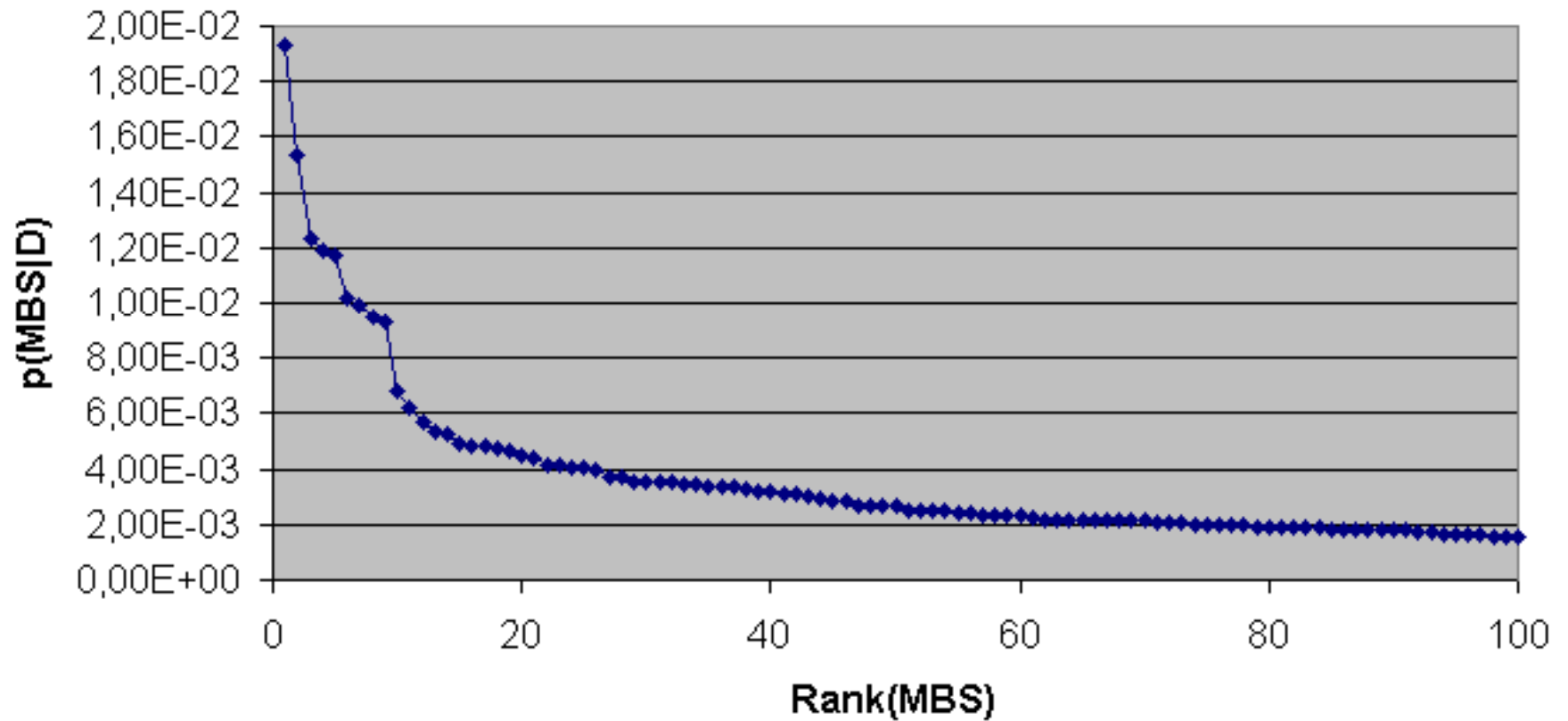


The peakness of the posteriors of the most probable MB sets and their MBM-based approximations. (46 variables, 1000 samples)

# Cumulative posterior of the most probable strongly relevant sets

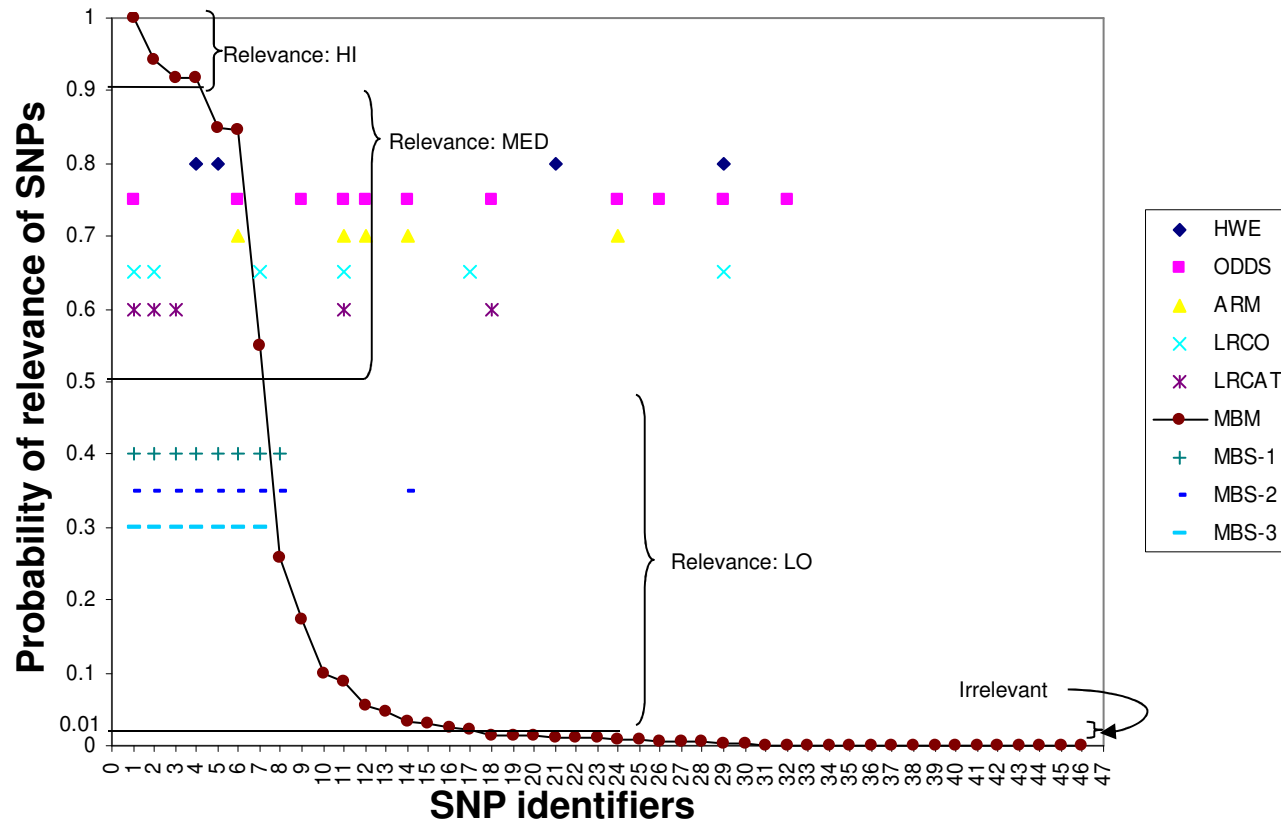


# MBS posteriors in Asthma



# Posteriors of strong relevance

*HWE* – Hardy-Weinberg equilibrium test, *ODDS* – odds ratio, *ARM* – Cochran-Armitage trend test, *LRCO* – logistic regression (continuous case), *LRCAT* – logistic regression (categorical case), *MBM* – Bayesian pairwise relevance, *MBS-1–9* relevant sets by Bayesian analysis. (only MBM values are numeric, others are arbitrary values for visualization)

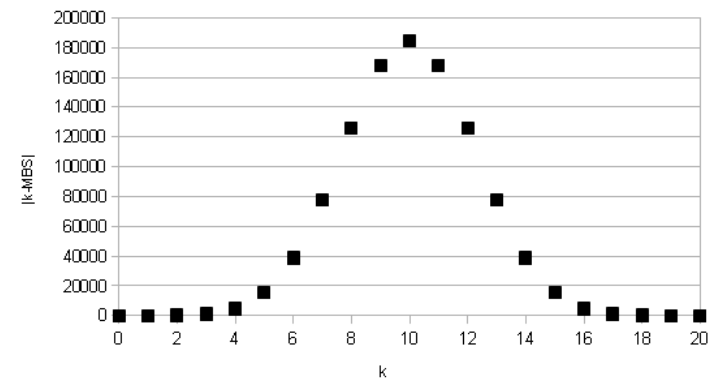
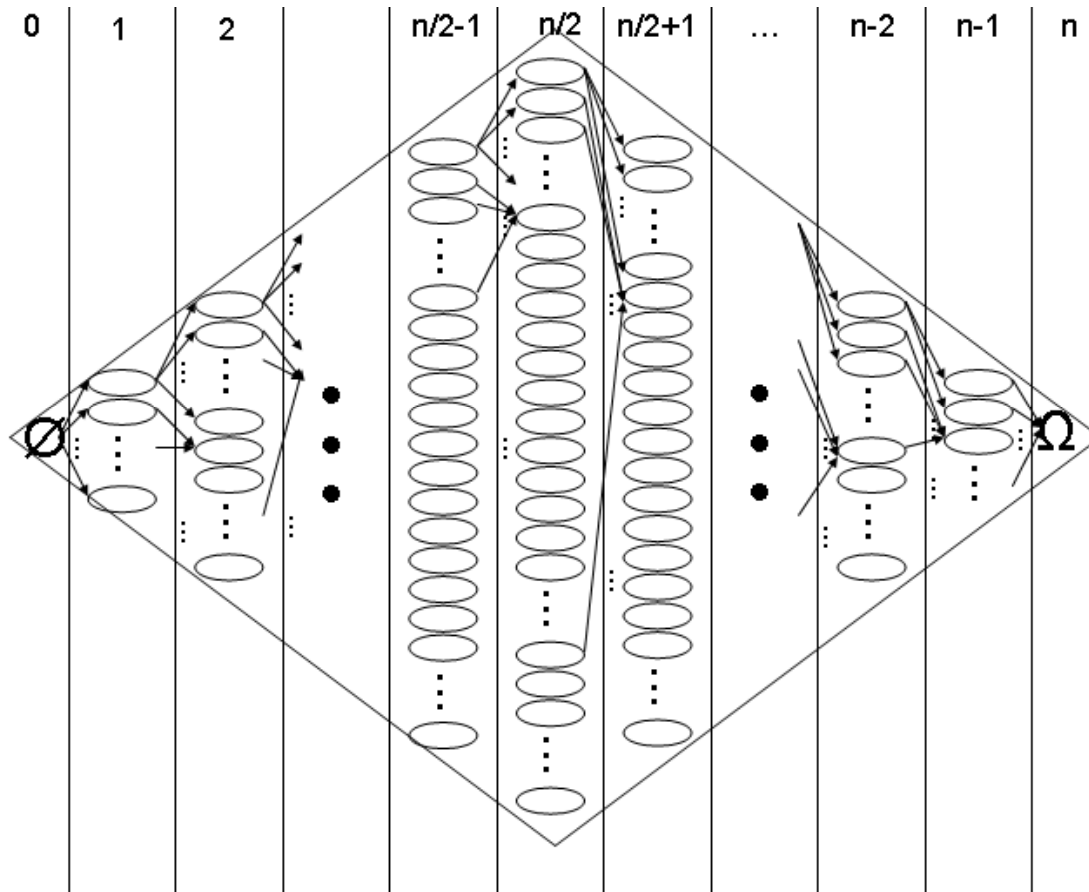


# Frequentist vs Bayesian statistics

Frequentist	Bayesian
-	Prior probabilities
Null hypothesis	-
Indirect: proving by refutation	Direct
Model selection	Model averaging
Likelihood ratio test	Bayes factor
p-value	-!
-!	Posterior probabilities
Confidence interval	Credible region
Significance level	Optimal decision based on Exp.Util.
Multiple testing problem	Remains, so → <b>complex model</b>
Model complexity dilemma	<b>Best achievable alternative</b>

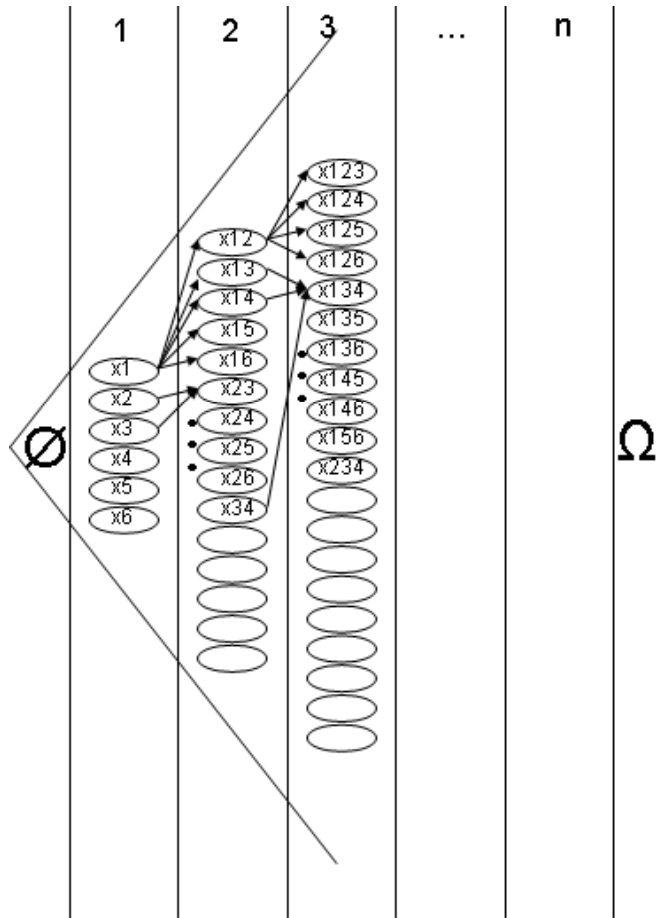
- Note: direct probabilistic statement!

# The subset space



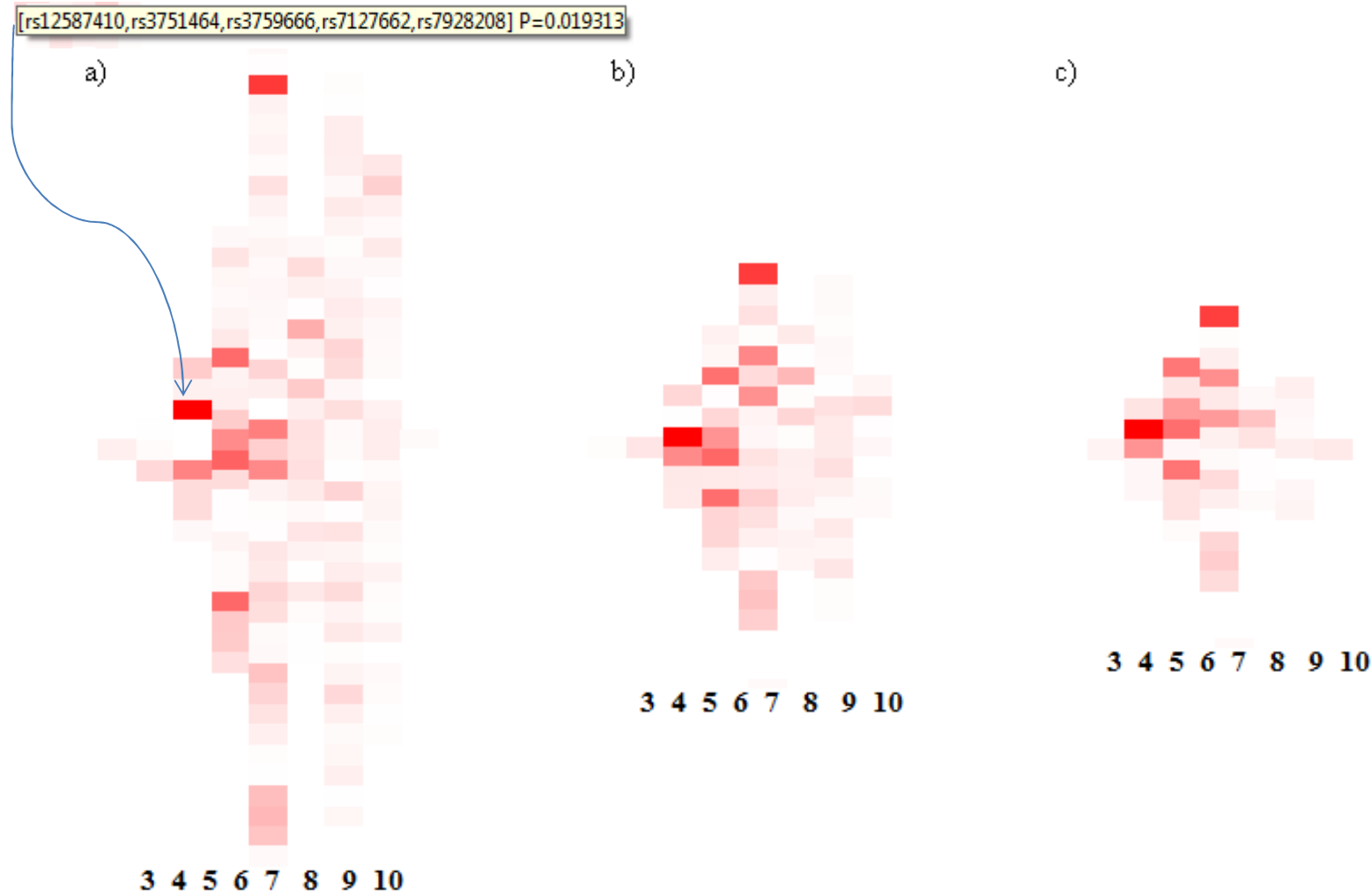


# The subset space II.



# An MBS heatmap in the subset space

[rs12587410,rs3751464,rs3759666,rs7127662,rs7928208] P=0.019313

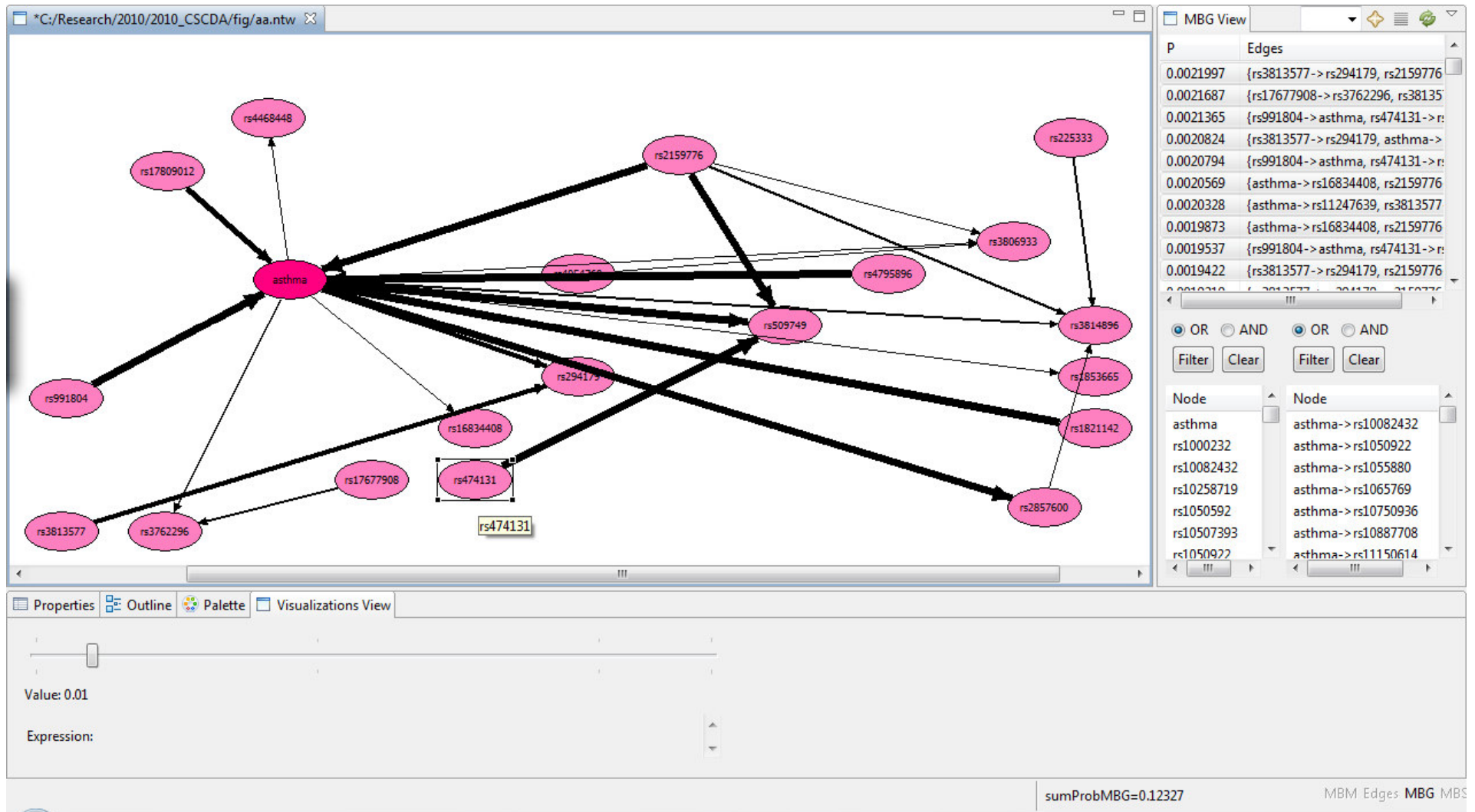


# Genagrid

- SGI Altix ICE
  - 5 TFLOPS
  - 1TB memory
  - 64x8 cores
  - FPGAs



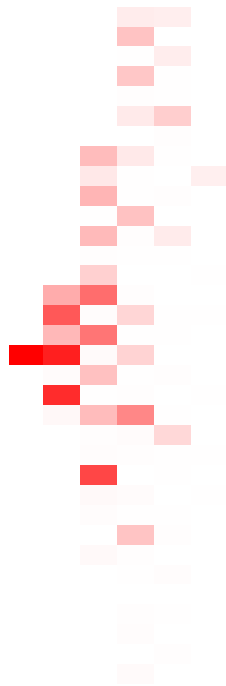
# BayesEye



# Marginal multivariate posteriors in the subset space?

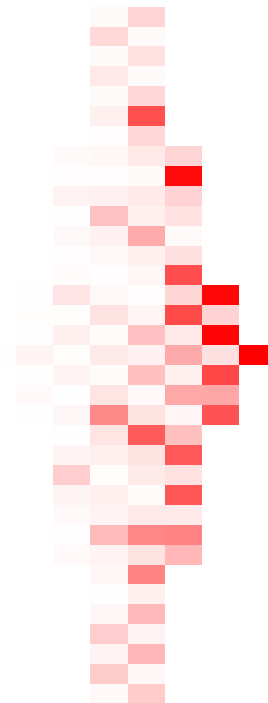
k-MBS-sub

$$p(G : s \subseteq \text{MBS}(G) \mid D_N)$$



k-MBS-sup

$$p(G : s \supseteq \text{MBS}(G) \mid D_N)$$



# Summary

- Feature relevance
- The feature subset selection problem
- Identification of biomarkers
  - Methods
- Challenges
  - Interpretation → Bayesian networks
  - Causality → Bayesian networks
  - Uncertainty → Bayesian statistics
- A Bayesian network based Bayesian approach to biomarker analysis