

GRAPHICAL MODELS FOR CAUSAL INFERENCE

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Introduction

Why do we need graphs?

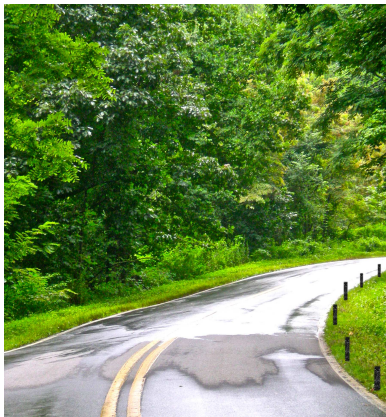


Figure: Motivating Example

Introduction

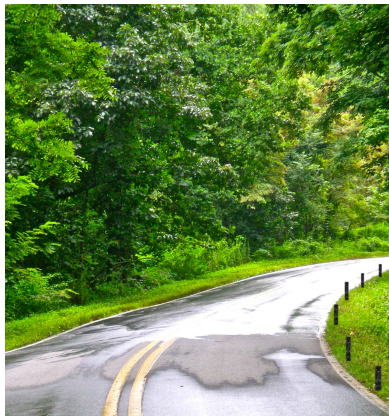
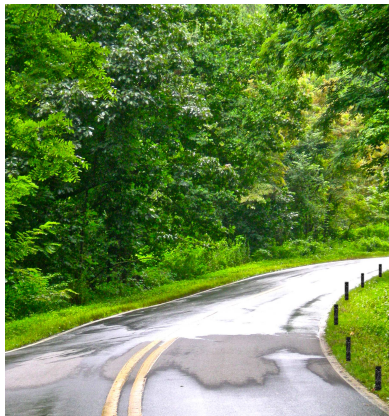


Figure: Motivating Example

Variables in the study:

- ▶ Season
- ▶ Sprinkler
- ▶ Rain
- ▶ Wetness of pavement(Wet)
- ▶ Slipperiness of pavement(Slippery)

Introduction



# Variables	Table size
5	32
6	64
7	128
8	256
9	512
10	1,024
20	1,048,576
30	1,073,741,824

Figure: Motivating Example

Introduction

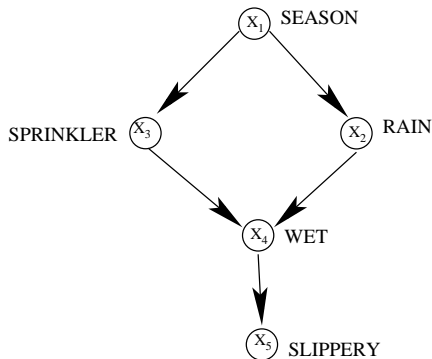


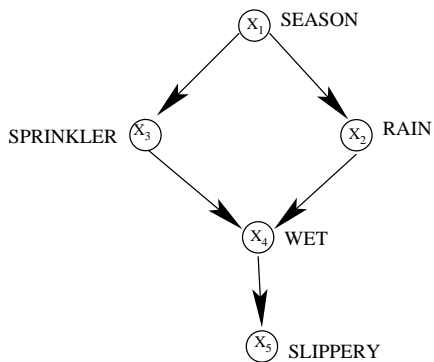
Figure: DAG Representation

Conditional Probability Distributions

- ▶ $P(X_1)$: 2
- ▶ $P(X_3|X_1)$: 4
- ▶ $P(X_2|X_1)$: 4
- ▶ $P(X_4|X_2, X_3)$: 8
- ▶ $P(X_5|X_4)$: 4

Total # of Table Entries = 22

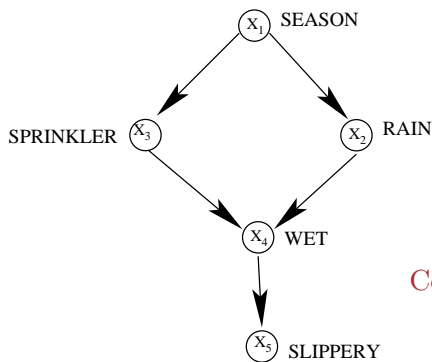
Graphs: Notations



- ▶ Adjacent Nodes
- ▶ Root and Leaf Nodes
- ▶ Skeleton
- ▶ Path
- ▶ Kinship Terminology

Figure: Bayesian Network representing dependencies

Graphs: Notations



Chain $X_1 \rightarrow X_3 \rightarrow X_4 \rightarrow X_5$
 $X_1 \rightarrow X_3 \rightarrow X_4 \rightarrow X_5$

Fork $X_3 \leftarrow X_1 \rightarrow X_2$

Collider $X_3 \rightarrow X_4 \leftarrow X_2$

Figure: Bayesian Network representing dependencies

Background Factors & Bi-directed Edges

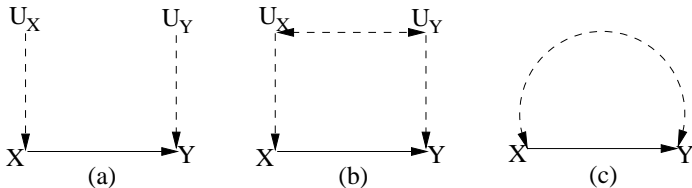


Figure: (a) Causal Model with background factors (b) & (c) Causal Model with correlated background factors

Decomposing joint distribution- $P(V)$

How would you decompose joint distribution $P(V)$ into smaller distributions?

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By applying Chain rule

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Let X_1, X_2, \dots, X_n be any arbitrary ordering of nodes in a DAG.
$$P(x_1, x_2, \dots, x_n) = \prod_j P(x_j | x_1, \dots, x_{j-1})$$

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Is it possible that conditional probability of some variable X_j is not sensitive to all its predecessors?

Decomposing joint distribution- $P(V)$

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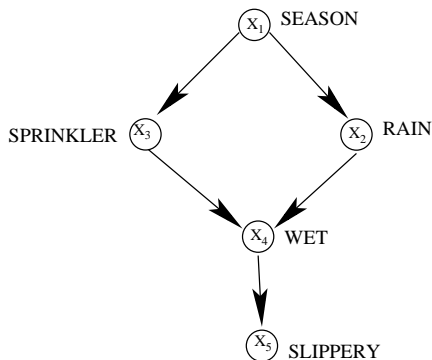
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Is it possible that conditional probability of some variable X_j is not sensitive to all its predecessors?

Yes!

Markovian Parents



Markovian Parents

$$X_1 : \phi$$

$$X_2 : \{X_1\}$$

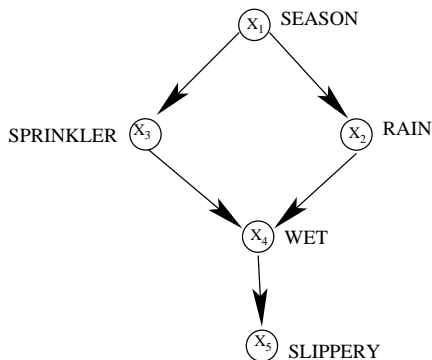
$$X_3 : \{X_1\}$$

$$X_4 : \{X_2, X_3\}$$

$$X_5 : \{X_4\}$$

Figure: Bayesian Network representing dependencies

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$$X_1 : \phi$$

$$X_2 : \{X_1\}$$

$$X_3 : \{X_1\}$$

$$X_4 : \{X_2, X_3\}$$

$$X_5 : \{X_4\}$$

Figure: Bayesian Network representing dependencies

$$P(x_1, x_2, x_3, x_4, x_5) = P(x_1)P(x_2|x_1)P(x_3|x_1)P(x_4|x_2, x_3)P(x_5|x_4)$$

Markov Compatibility

Let $V = \{x_1, x_2, \dots, x_n\}$ be the set of observed nodes and pa_i be the Markovian parents of x_i . Then,

$$P(v) = P(x_1, x_2, \dots, x_n) = \prod_i P(x_i | pa_i).$$

Markov Compatibility

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Definition (Markov Compatibility)

If a probability distribution P admits Markovian factorization of observed nodes relative to DAG G , we say that G and P are Markov compatible.

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Definition (Markov Compatibility)

If a probability distribution P admits Markovian factorization of observed nodes relative to DAG G , we say that G and P are Markov compatible.

Example

X	Y	$P(X, Y)$
1	1	0.225
1	0	0.375
0	1	0.125
0	0	0.275

Markov Compatible DAGs:

$X \rightarrow Y$

$X \leftarrow Y$

Testing Markov Compatibility

Given a DAG G and distribution P , how can you conclude that P and G are compatible?

Testing Markov Compatibility

Given a DAG G and distribution P , how can you conclude that P and G are compatible?

- ▶ Parents shielding tests
 - ▶ non-descendants
 - ▶ predecessors
- ▶ d-separation

d-Separation

Definition

Let X , Y and Z be disjoint sets in DAG G . X and Y are d-separated by Z (written $(X \perp\!\!\!\perp Y|Z)_G$) if and only if Z blocks every path from a node in X to a node in Y .

d-Separation

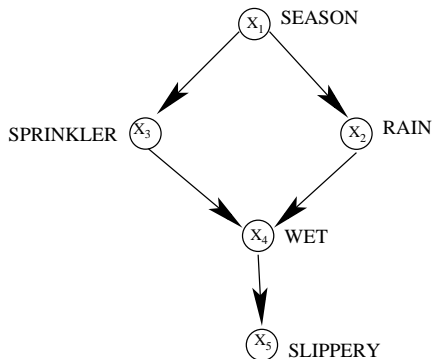
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A path p is said to be d-separated(or blocked) by a set of nodes Z if and only if :

- (1) p contains a chain $i \rightarrow m \rightarrow j$ or a fork $i \leftarrow m \rightarrow j$ such that the middle node m is in Z , or
- (2) ...

Example: d-Separation



- ▶ $X_1 \perp\!\!\!\perp X_4 \mid X_2, X_3$
- ▶ $X_3 \perp\!\!\!\perp X_5 \mid X_4$
- ▶ $X_1 \perp\!\!\!\perp X_5 \mid X_4$

d-Separation

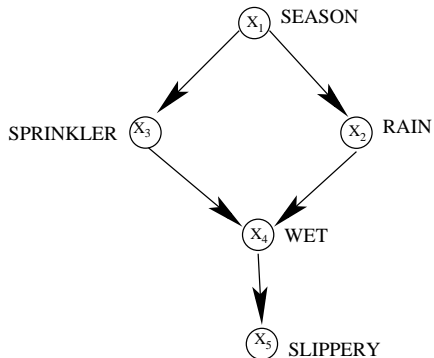
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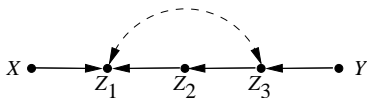
- (1) p contains a chain $i \rightarrow m \rightarrow j$ or a fork $i \leftarrow m \rightarrow j$ such that the middle node m is in Z , or
- (2) p contains an inverted fork (or collider) $i \rightarrow m \leftarrow$ such that the middle node m is not in Z and such that no descendant of m is in Z .

Example: d-Separation

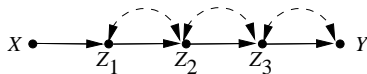


► $X_3 \perp\!\!\!\perp X_2 | X_1$

Example: d-Separation

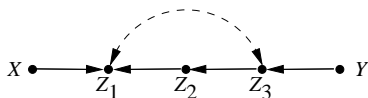


(a)

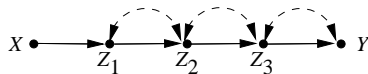


(b)

Example: d-Separation



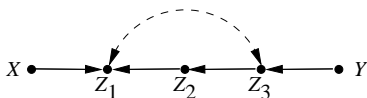
(a)



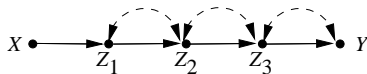
(b)

Case (a): $X \perp\!\!\!\perp Y \mid \phi$

Example: d-Separation



(a)



(b)

Case (a): $X \perp\!\!\!\perp Y \mid \phi$

Case (b): $X \not\perp\!\!\!\perp Y$

d-Separation

When is it impossible to d-separate 2 non-adjacent nodes X and Y ?

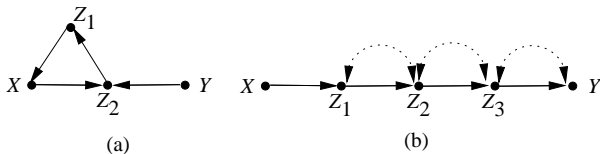
Do we need to test all sets for possible separation?

Inducing path

Definition

Path between 2 nodes X and Y is termed inducing if every non-terminal node on the path:

- (i) is a collider and
- (ii) an ancestor of either X or Y (or both)



Note: There are no separators for X and Y .

The Five Necessary Steps of Causal Analysis

- Define** Express the target quantity Q as property of the model M .
- Assume** Express causal assumptions in structural or graphical form.
- Identify** Determine if Q is identifiable.
- Estimate** Estimate Q if it is identifiable; approximate it, if it is not.
- Test** If M has testable implications

A “Mini” Turing Test in Causal Conversation

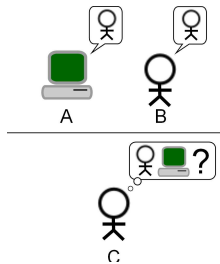
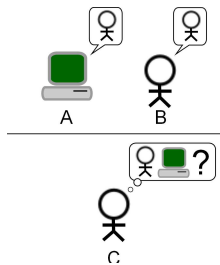


Figure: Turing Test

A “Mini” Turing Test in Causal Conversation



Input: Story

Question: What if? What is? Why?

Answers: I believe that...

Figure: Turing Test

A “Mini” Turing Test in Causal Conversation

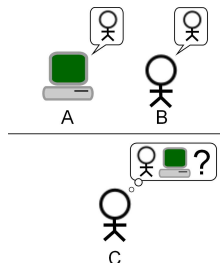
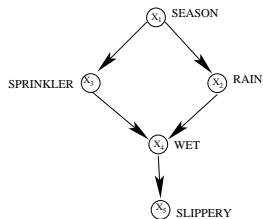
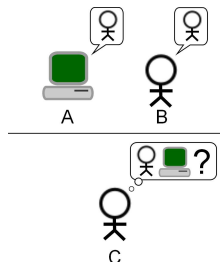


Figure: Turing Test

The Story



A “Mini” Turing Test in Causal Conversation



The Story

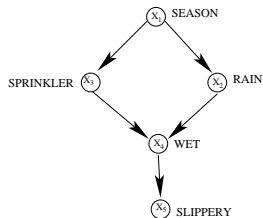
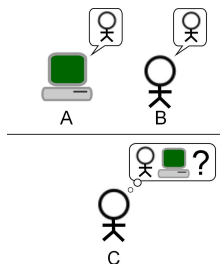


Figure: Turing Test

Q1: If the season is dry and the pavement is slippery, did it rain?

A “Mini” Turing Test in Causal Conversation



The Story

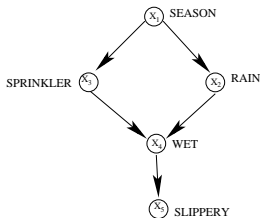
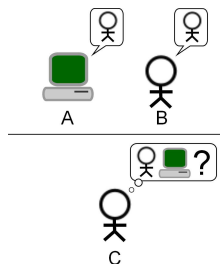


Figure: Turing Test

Q1: If the season is dry and the pavement is slippery, did it rain?

A1: Unlikely, it is more likely that the sprinkler was ON with a very slight possibility that it is not even wet.

A “Mini” Turing Test in Causal Conversation



The Story

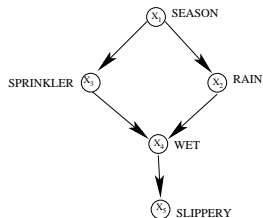


Figure: Turing Test

Q2: But what if we see that the sprinkler is OFF?

A “Mini” Turing Test in Causal Conversation

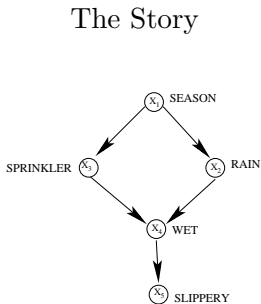
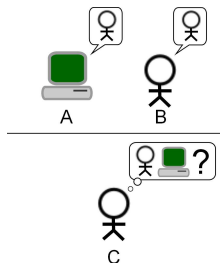


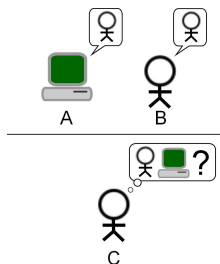
Figure: Turing Test

Q2: But what if we see that the sprinkler is OFF?

A2: Then it is more likely that it rained.

Without graphs, # of Table Entries = 32

A “Mini” Turing Test in Causal Conversation



The Story

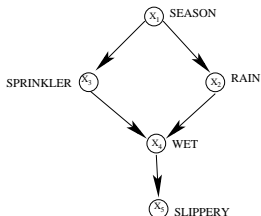
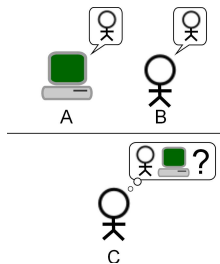


Figure: Turing Test

Q3: Do you mean that if we actually turn the sprinkler ON, the rain will be less likely?

A “Mini” Turing Test in Causal Conversation



The Story

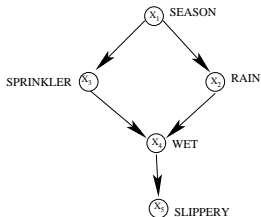


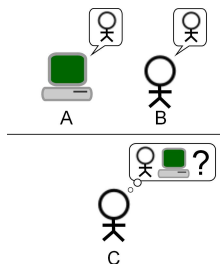
Figure: Turing Test

Q3: Do you mean that if we actually turn the sprinkler ON, the rain will be less likely?

A3: No, the likelihood of rain would remain the same but the pavement would surely get wet.

Without graphs, # of Table Entries = $32 * 32$

A “Mini” Turing Test in Causal Conversation



The Story

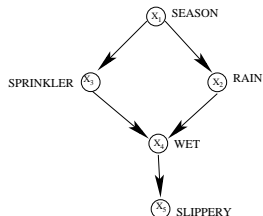
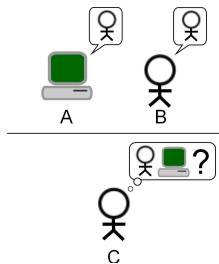


Figure: Turing Test

Q4: Suppose we see that the sprinkler is ON and pavement is wet. What if the sprinkler were OFF?

A “Mini” Turing Test in Causal Conversation



The Story

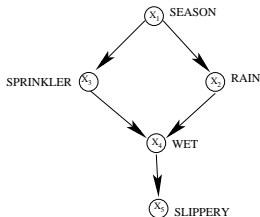


Figure: Turing Test

Q4: Suppose we see that the sprinkler is ON and pavement is wet. What if the sprinkler were OFF?

A4: The pavement would be dry because the season is likely to be dry

Without graphs, what would be the # of table entries?

Interventions

Query: Would the pavement be slippery if we *make sure that* the the sprinkler is on?

Compute: $P(x_5 | do(x_3))$

May be equivalently represented as:

(a) $P(x_5 | \hat{x}_3)$

(b) $P_{x_3}(x_5)$

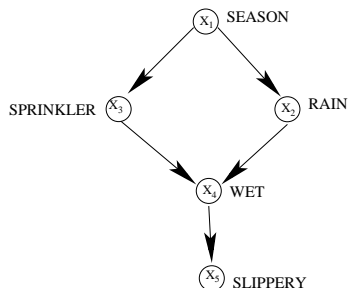


Figure: DAG before intervention

Interventions

Compute: $P(x_5 | do(x_3))$

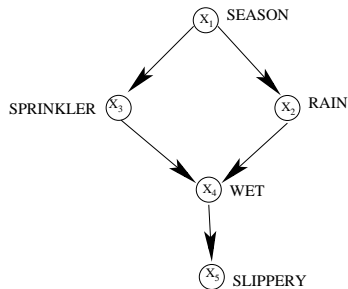


Figure: DAG before intervention

$$P(v) = P(x_1)P(x_2|x_1)P(x_3|x_1)P(x_4|x_2, x_3)P(x_5|x_4)$$

Interventions

Compute: $P(x_5 | do(x_3))$

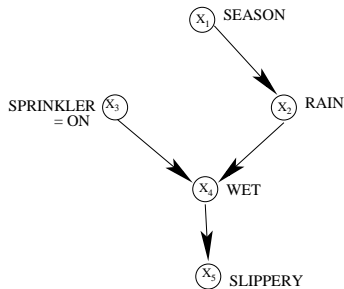
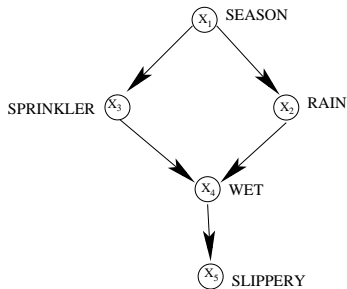


Figure: DAG before intervention

Figure: DAG after intervention

$$P(v) = P(x_1)P(x_2|x_1) \frac{P(x_3|x_1)}{P(x_3)} P(x_4|x_2, x_3)P(x_5|x_4)$$

$$P(x_1, x_2, x_4, x_5 | do(x_3)) = P(x_1)P(x_2|x_1)P(x_4|x_2, x_3)P(x_5|x_4)$$

Interventions

Compute: $P(x_5 | do(x_3))$

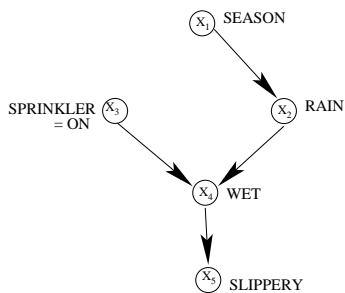
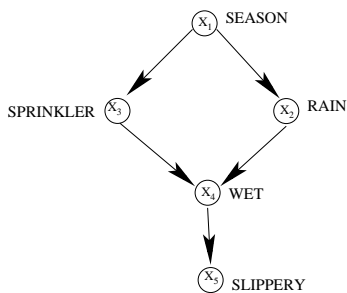


Figure: DAG before intervention

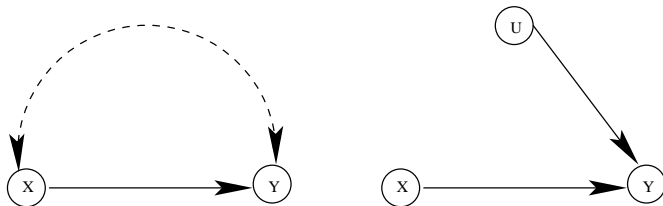
Figure: DAG after intervention

$$P(v) = P(x_1)P(x_2|x_1) \frac{P(x_3|x_1)}{P(x_3|x_1)} P(x_4|x_2, x_3)P(x_5|x_4)$$

$$P(x_5 | do(x_3)) = \sum_{x_1, x_2, x_4} P(x_1)P(x_2|x_1)P(x_4|x_2, x_3)P(x_5|x_4)$$

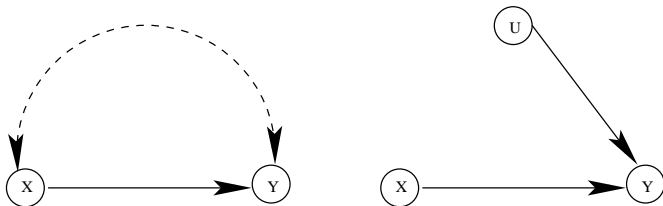
Note: $P(x_5 | do(x_3)) \neq P(x_5 | x_3)$ i.e. Doing \neq Seeing

Examples



Question : Can you estimate $P(y|do(x))$, given $P(x, y)$?

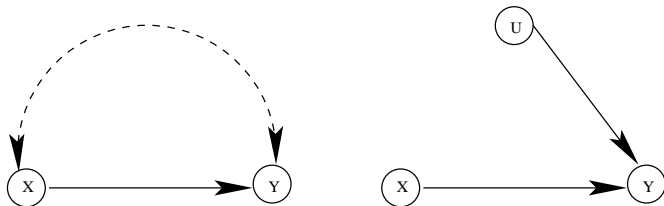
Examples



Question : Can you estimate $P(y|do(x))$, given $P(x, y)$?

NO!

Examples



Question : Can you estimate $P(y|do(x))$, given $P(x, y)$?

NO!

$$P(x, y) = \sum_u P(x, y, u) = \sum_u P(y|x, u)P(x|u)P(u)$$

$$P(y|do(x)) = \sum_u P(y|x, u)P(u)$$

Identifiability

Definition

Let $Q(M)$ be any computable quantity of a model M .

Identifiability

Definition

Let $Q(M)$ be any computable quantity of a model M . We say that Q is identifiable in a class M of models if, for any pairs of models M_1 and M_2 from M , $Q(M_1) = Q(M_2)$ whenever $P_{M_1}(v) = P_{M_2}(v)$.

Estimating causal effect

Adjustment for direct causes

Compute: $P(y|\hat{x})$

$$P(x, y, z, w) = P(y|x, w)P(x|z)P(w|z)P(z)$$

$$P(y, z, w|do(x)) = P(y|x, w)P(w|z)P(z) \frac{P(x|z)}{P(x|z)}$$

$$P(y, z, w|do(x)) = \frac{P(x, y, z, w)}{P(x|z)}$$

$$P(y|do(x)) = \sum_{z, w} P(yw|x, z)P(z) = \sum_z P(y|x, z)P(z)$$

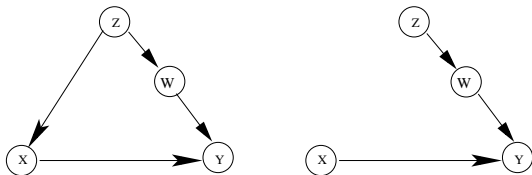


Figure: DAGs before and after intervention

Theorem (Adjustment for direct causes)

Let PA_i denote the set of direct causes of X_i and let Y be any set of variables disjoint of $\{X_i \cup PA_i\}$. The causal effect of X_i on Y is given by:

$$P(y|\hat{x}_i) = \sum_{pa_i} P(y|x_i, pa_i)P(pa_i)$$

where $P(y|x_i, pa_i)$ and $P(pa_i)$ represent pre-interventional probabilities.

Example: Adjustment for direct causes

Query: Would the pavement be slippery if we *make sure that* the sprinkler is on?

$$P(x_5 | \hat{x}_3) = \sum_{x_1} P(x_5 | x_3, x_1) P(x_1)$$

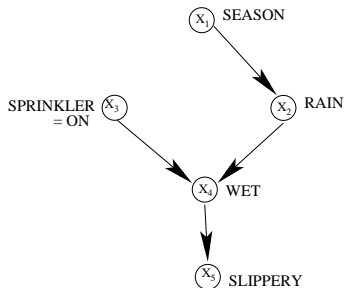
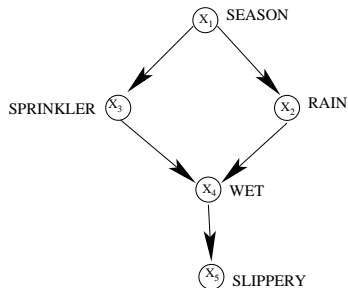


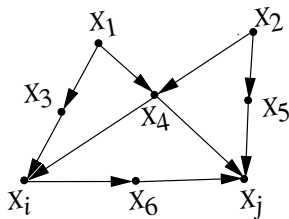
Figure: DAG before intervention

Figure: DAG after intervention

Estimating Causal Effect

Compute: $P(X_j | do(X_i))$

How can we find a set Z of concomitants that are sufficient for identifying causal effect?



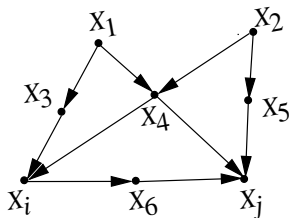
Back-door Criterion for Identifiability

Definition (Pearl-1993)

A set of variables Z satisfies the back-door criterion relative to an ordered pair of variables (X_i, X_j) in a DAG G if:

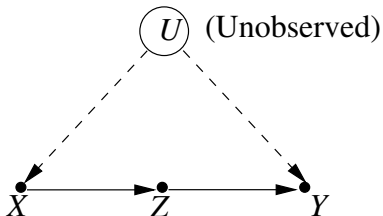
- (i) no node in Z is a descendant of X_i ; and
- (ii) Z blocks every path between X_i and X_j that contains an arrow into X_i .

$$P(x_j | do(x_i)) = \sum_z P(x_j | x_i, z) P(z)$$



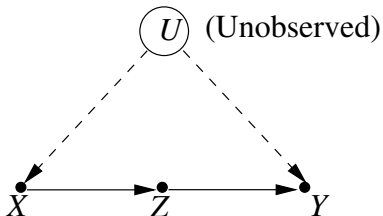
Estimating causal effect: $P(y|do(x))$

- ▶ Can you adjust for direct cause?



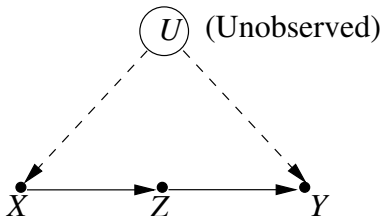
Estimating causal effect: $P(y|do(x))$

- ▶ Can you adjust for direct cause? **NO!**



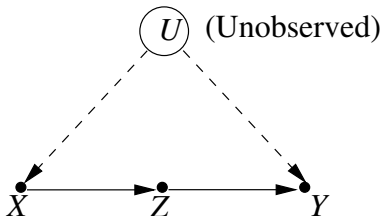
Estimating causal effect: $P(y|do(x))$

- ▶ Can you apply backdoor criterion?



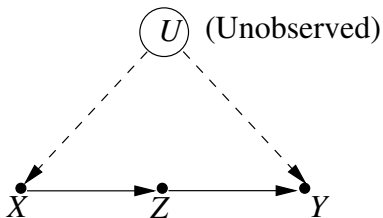
Estimating causal effect: $P(y|do(x))$

- ▶ Can you apply backdoor criterion? **NO!**



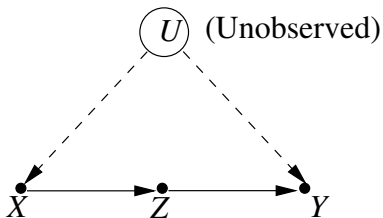
Estimating causal effect: $P(y|do(x))$

- ▶ Is $P(y|do(x))$ identifiable?



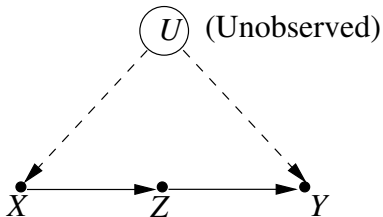
Estimating causal effect: $P(y|do(x))$

- ▶ Is $P(y|do(x))$ identifiable? **YES!**



Estimating causal effect: $P(y|do(x))$

Given: $P(y|\hat{x}) = \sum_z P(y|\hat{z})P(z|\hat{x})$



Estimating causal effect: $P(y|do(x))$

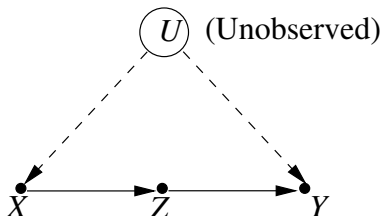
$$\text{Given: } P(y|\hat{x}) = \sum_z P(y|\hat{z})P(z|\hat{x})$$

$$P(z|\hat{x}) = P(z|x)$$

$$P(y|\hat{z}) = \sum_{x'} P(y|x', z)P(x')$$

Therefore,

$$P(y|\hat{x}) = \sum_z P(z|x) \sum_{x'} P(y|x', z)P(x')$$



Front-door Criterion for Identifiability

Definition (Pearl-1995)

A set of variables Z satisfies the front-door criterion relative to an ordered pair of variables (X_i, X_j) in a DAG G if:

- (i) Z intercepts all directed paths from X to Y ; and
- (ii) there is no unblocked back-door path from X to Z ; and
- (iii) all back-door paths from Z to Y are blocked by X

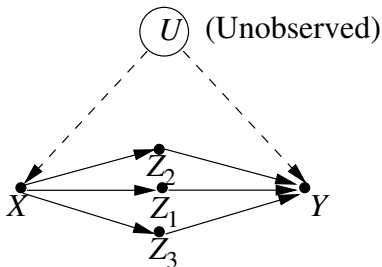


Figure: Frontdoor criterion is satisfied by $Z = \{Z_1, Z_2, Z_3\}$

Front-door Adjustment

If Z satisfies the front door criterion relative to (X, Y) and if $P(x, z) > 0$, then the causal effect of X on Y is identifiable and is given by:

$$P(y|\hat{x}) = \sum_z P(z|x) \sum_{x'} P(y|x', z)P(x')$$

Estimating causal effect $P(y|do(x))$

How can you syntactically derive claims about interventions?

Estimating causal effect $P(y|do(x))$

How can you syntactically derive claims about interventions?

- ▶ do-calculus

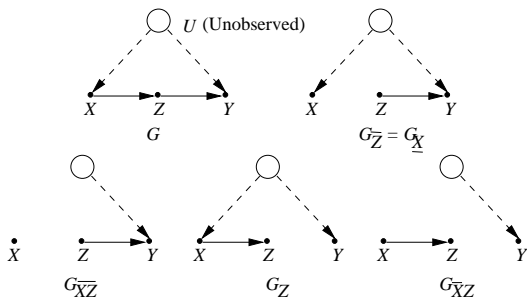
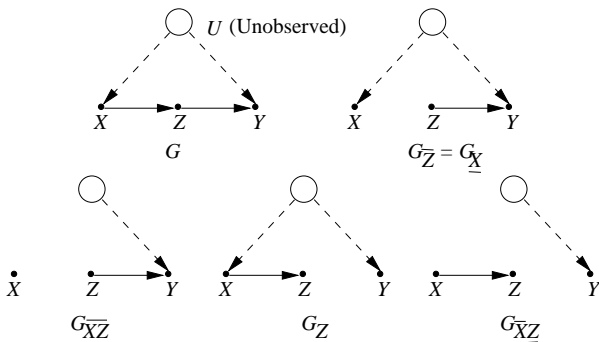


Figure: Subgraphs of G used in the derivation of causal effects.

do-Calculus-[Pearl-1995]

Rule-1 Insertion or deletion of observations

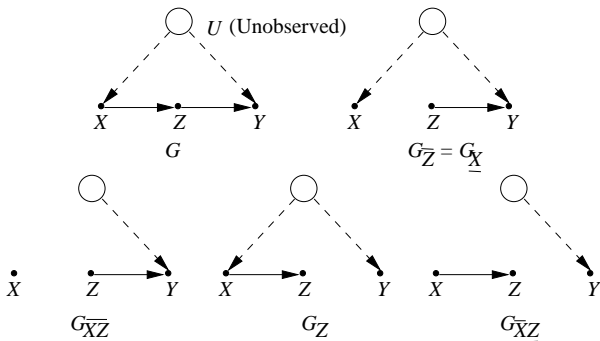
$$P(y|\hat{x}, z, w) = P(y|\hat{x}, w) \text{ if } (Y \perp\!\!\!\perp Z|X, W)_{G_{\bar{X}}}$$



do-Calculus-[Pearl-1995]

Rule-2 Action/Observation exchange

$$P(y|\hat{x}, \hat{z}, w) = P(y|\hat{x}, z, w) \text{ if } (Y \perp\!\!\!\perp Z|X, W)_{G_{\overline{XZ}}}$$

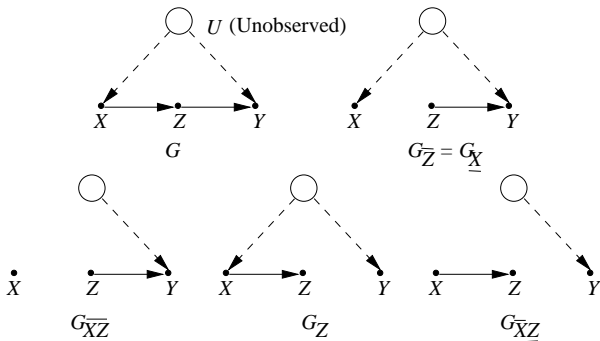


do-Calculus-[Pearl-1995]

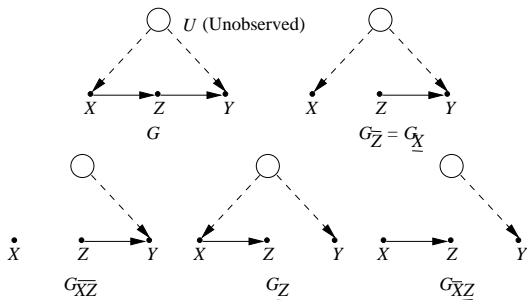
Rule-3 Insertion or deletion of actions

$$P(y|\hat{x}, \hat{z}, w) = P(y|\hat{x}, w) \text{ if } (Y \perp\!\!\!\perp Z|X, W)_{G_{\overline{X}, \overline{Z(W)}}}$$

where $Z(W)$ is the set of Z nodes that are not ancestors of any W node in $G_{\overline{X}}$



Deriving causal effect using do-calculus



Compute: $P(y|\hat{z})$

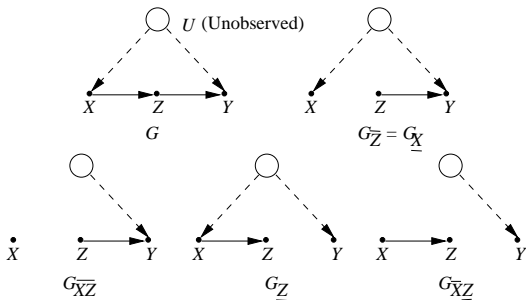
$$P(y|\hat{z}) = \sum_x P(y|x, \hat{z})P(x|\hat{z})$$

$$P(x|\hat{z}) = P(x) \text{ since } (Z \perp\!\!\!\perp X)_{G_{\bar{Z}}}$$

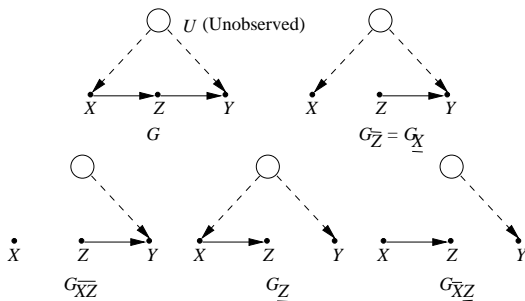
$$P(y|x, \hat{z}) = P(y|x, z) \text{ since } (Z \perp\!\!\!\perp Y|X)_{G_{\underline{Z}}}$$

$$P(y|\hat{z}) = \sum_x P(y|x, z)P(x)$$

Prove: $P(y|\hat{x}) = \sum_z P(y|\hat{z})P(z|\hat{x})$

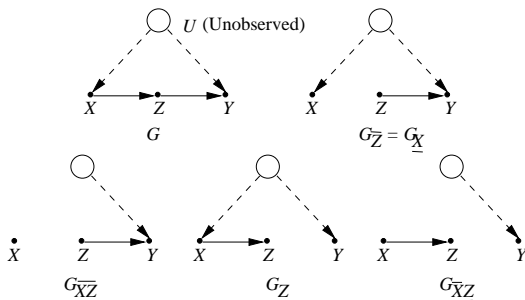


Prove: $P(y|\hat{x}) = \sum_z P(y|\hat{z})P(z|\hat{x})$



$$P(y|\hat{x}) = \sum_z P(yz|\hat{x}) = \sum_z P(y|\hat{x}z)P(z|\hat{x})$$

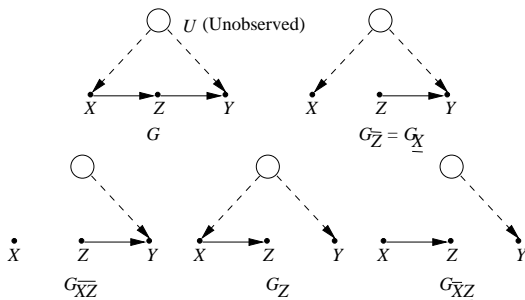
Prove: $P(y|\hat{x}) = \sum_z P(y|\hat{z})P(z|\hat{x})$



$$P(y|\hat{x}) = \sum_z P(yz|\hat{x}) = \sum_z P(y|\hat{x}z)P(z|\hat{x})$$

$$P(y|\hat{x}z) = P(y|\hat{z}\hat{x}) \text{ since } Y \perp\!\!\!\perp Z \text{ in } G_{\bar{X}\underline{Z}}$$

Prove: $P(y|\hat{x}) = \sum_z P(y|\hat{z})P(z|\hat{x})$



$$P(y|\hat{x}) = \sum_z P(yz|\hat{x}) = \sum_z P(y|\hat{x}z)P(z|\hat{x})$$

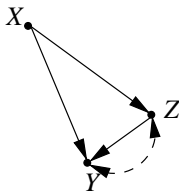
$$P(y|\hat{x}z) = P(y|\hat{z}\hat{x}) \text{ since } Y \perp\!\!\!\perp Z \text{ in } G_{\bar{X}\bar{Z}}$$

$$= P(y|\hat{z}) \text{ since } Y \perp\!\!\!\perp X \text{ in } G_{\bar{X}\underline{Z}}$$

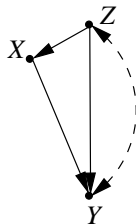
Graphical Models in which $P(y|\hat{x})$ is Identifiable



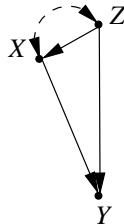
(a)



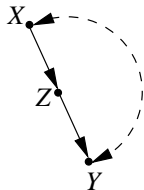
(b)



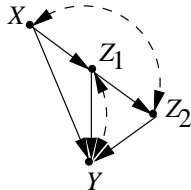
(c)



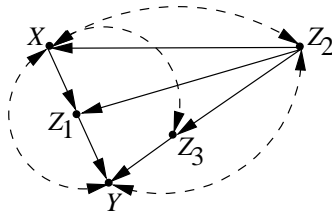
(d)



(e)

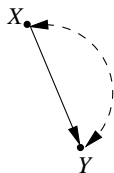


(f)

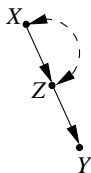


(g)

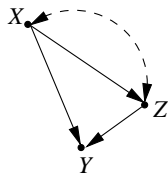
Graphical Models in which $P(y|\hat{x})$ is not Identifiable



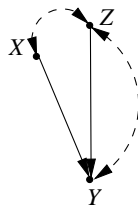
(a)



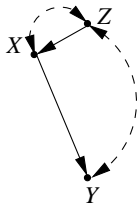
(b)



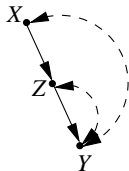
(c)



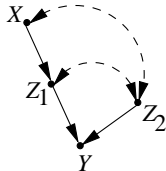
(d)



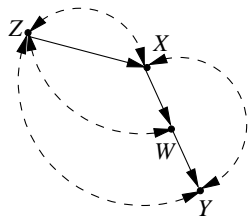
(e)



(f)



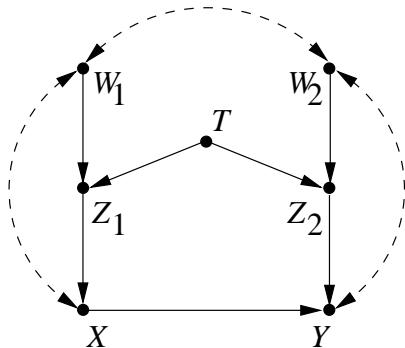
(g)



(h)

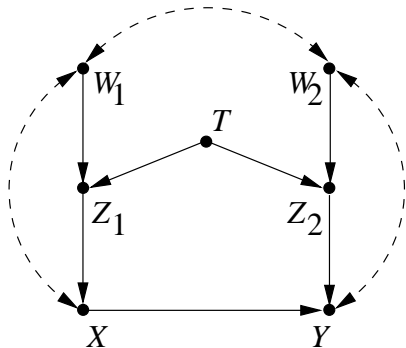
C-components and C-factor

Two variables are said to be in the same C-component if they are connected by a path comprising of only bi-directional edges [Tian & Pearl, 2002].



C-components and C-factor

Two variables are said to be in the same C-component if they are connected by a path comprising of only bi-directional edges [Tian & Pearl, 2002].



$$S_1 = \{X, Y, W_1, W_2\}$$

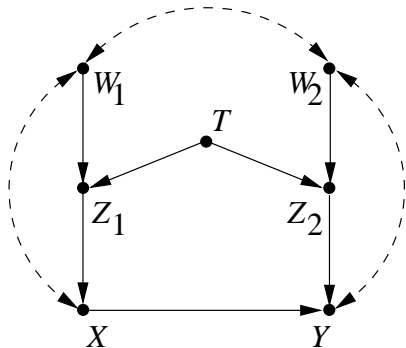
$$S_2 = \{Z_1\}$$

$$S_3 = \{Z_2\}$$

$$S_4 = \{T\}$$

C-components and C-factor

Two variables are said to be in the same C-component if they are connected by a path comprising of only bi-directional edges [Tian & Pearl, 2002].



$$S_1 = \{X, Y, W_1, W_2\}$$

$$S_2 = \{Z_1\}$$

$$S_3 = \{Z_2\}$$

$$S_4 = \{T\}$$

C-factor: $Q[S_i](v) = P_{v \setminus s_i}(s_i)$

Identifiability of C-factor

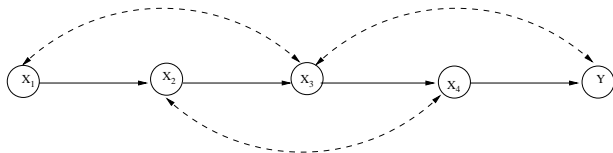
Lemma (Tian & Pearl, 2002)

Let a topological order over V be $V_1 < V_2 < \dots < V_n$ and let $V^{(i)} = \{V_1, V_2, \dots, V_i\}$, $i = 1, \dots, n$ and $V^{(0)} = \emptyset$. For any set C , let G_C denote the subgraph of G composed only of variables in C . Then:

(i) Each C-factor Q_j , $j = 1, \dots, k$ is identifiable and is given by:

$$Q_j = \prod_{\{i: V_i \in S_j\}} P(v_i | v^{(i-1)})$$

Example: Identifiability of C-factor



Admissible order: $X_1 < X_2 < X_3 < X_4 < Y$

$$Q_1 = P(x_4|x_1, x_2, x_3)P(x_2|x_1)$$

$$Q_2 = P(y|x_1, x_2, x_3, x_4)P(x_3|x_1, x_2)P(x_1)$$

Necessary and Sufficient condition for identifiability of $P_x(v)$

Theorem (Tian & Pearl, 2002)

Let X be a singleton. $P_x(v)$ is identifiable if and only if there is no bi-directed path connecting X to any of its children.

Necessary and Sufficient condition for identifiability of $P_x(v)$

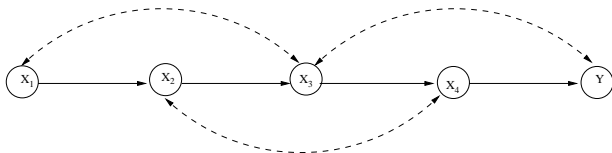
Theorem (Tian & Pearl, 2002)

Let X be a singleton. $P_x(v)$ is identifiable if and only if there is no bi-directed path connecting X to any of its children. When $P_x(v)$ is identifiable, it is given by:

$$P_x(v) = \frac{P(v)}{Q^X} \sum_x Q^X,$$

where Q^X is the c-factor corresponding to the c-component S^X that contains X .

Example: Necessary and Sufficient condition for identifiability of $P_x(v)$



Admissible order: $X_1 < X_2 < X_3 < X_4 < Y$

$$Q_1 = P(x_4|x_1, x_2, x_3)P(x_2|x_1)$$

$$Q_2 = P(y|x_1, x_2, x_3, x_4)P(x_3|x_1, x_2)P(x_1)$$

$$\begin{aligned} P_{x_1}(x_2, x_3, x_4, y) &= Q_1 \sum_{x_1} Q_2 \\ &= P(x_4|x_1, x_2, x_3)P(x_2|x_1) \\ &\quad \sum_{x_1} P(y|x_1, x_2, x_3, x_4)P(x_3|x_1, x_2)P(x_1) \end{aligned}$$

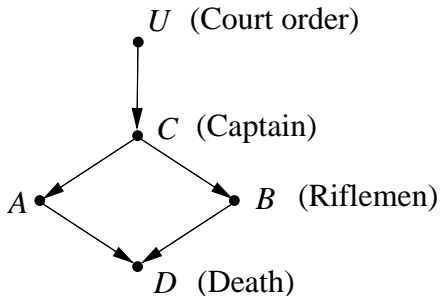
Causal Effect Identifiability

Identification of $P_x(y|z)$ where $X \cap Y \cap Z = \phi$ and X is not necessarily a singleton, [Shpitser & Pearl,2006]

- ▶ Hedge Criterion
- ▶ IDC - Sound and Complete Algorithm

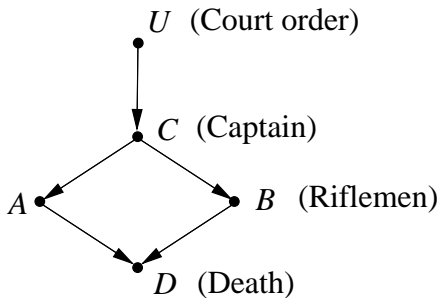
Counterfactuals

Query: Would the prisoner be dead had rifleman A not shot him, given that the prisoner is dead and rifleman A shot him?



Counterfactuals

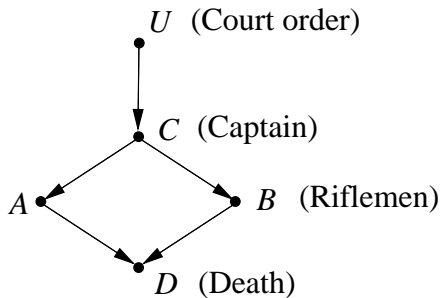
Query: Would the prisoner be dead had rifleman A not shot him, given that the prisoner is dead and rifleman A shot him?



► Abduction

Counterfactuals

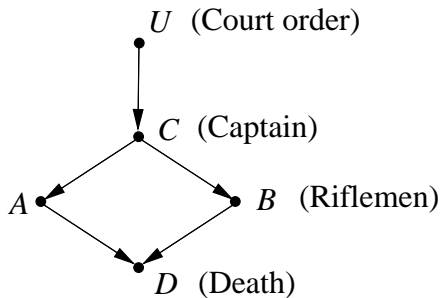
Query: Would the prisoner be dead had rifleman A not shot him, given that the prisoner is dead and rifleman A shot him?



- ▶ Abduction
- ▶ Intervention

Counterfactuals

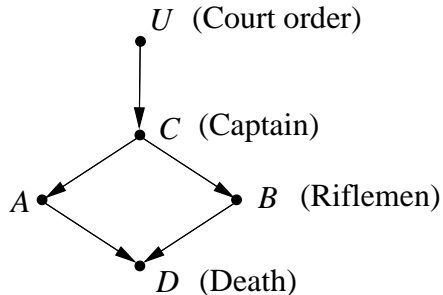
Query: Would the prisoner be dead had rifleman A not shot him, given that the prisoner is dead and rifleman A shot him?



- ▶ Abduction
- ▶ Intervention
- ▶ Prediction

Counterfactuals

Query: Would the prisoner be dead had rifleman A not shot him, given that the prisoner is dead and rifleman A shot him?



Model M

$$C = U$$

$$A = C$$

$$B = C$$

$$D = A \vee B$$

Facts: D

Conclusions:

$$U, A, B, C, D$$

Model $M_{\neg A}$

$$C = U$$

$$\neg A$$

$$B = C$$

$$D = A \vee B$$

Facts: U

Conclusions:

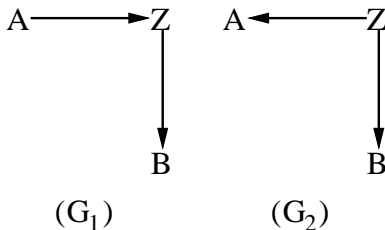
$$U, \neg A, B, C, D$$

Markov Equivalence

Given 2 models, is there a test that would tell them apart?

Definition

Two graphs G_1 and G_2 are said to be Markov equivalent if every d-separation condition in one also holds in the other. .



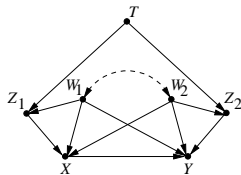
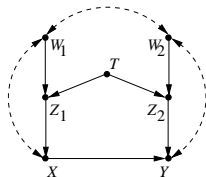
Markov Equivalence

Given 2 models, is there a test that would tell them apart?

Definition

Two graphs G_1 and G_2 are said to be Markov equivalent if every d-separation condition in one also holds in the other.

Are these DAGs Markov Equivalent?



Hard to enumerate all separation conditions.

Observational Equivalence

Theorem (Verma & Pearl 1990)

Two DAGs are observationally equivalent iff they have the same sets of edges and the same sets of v -structures, that is, two converging arrows whose tails are not connected by an arrow.

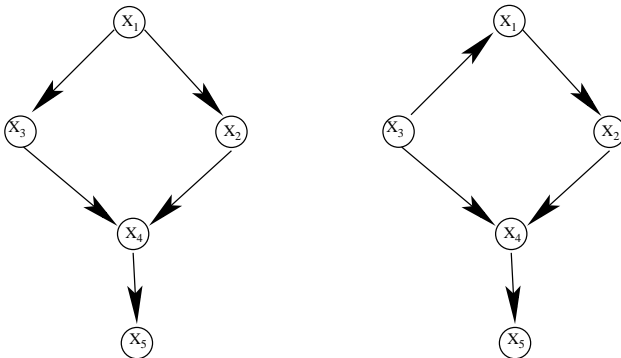


Figure: Observationally Equivalent DAGs

Markov Equivalence and Observational Equivalence

If two DAGs are Markov Equivalent, then they are Observationally Equivalent as well. True/False?

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True if all variables are observed(i.e. no bi-directed edges) and False otherwise.

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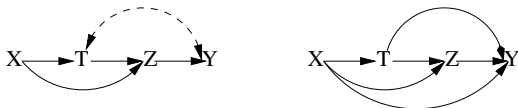


Figure: DAGs that are Markov Equivalent but not Observationally Equivalent

How would you distinguish between the two?

- ▶ Verma Constraints (Refer slide:115)

Ancestral Graphs

Definition (Ancestral Graphs)

A graph which may contain directed or bi-directed edges is ancestral if:

- (i) there are no directed cycles
- (ii) whenever there is an edge $X \longleftrightarrow Y$, then there is no directed path from X to Y or from Y to X .

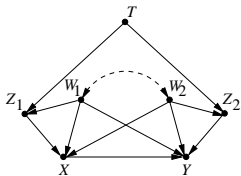


Figure: Ancestral graph

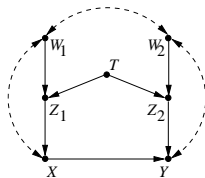


Figure: Not an Ancestral graph

Maximal Ancestral Graphs (MAGs)

Definition (Spirtes & Richardson, 2002)

An ancestral graph is said to be maximal if, for every pair of non-adjacent nodes X, Y there exists a set Z such that X and Y are d-separated conditional on Z .



Figure: DAG and its corresponding MAG

Construction of a MAG

Given : DAG G

Step-1: Construct a graph M comprising of:

- (i) all nodes in G
- (ii) all uni-directional edges in G

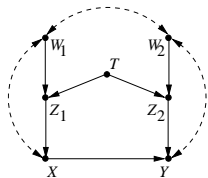


Figure: DAG G

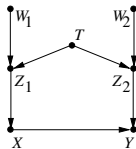


Figure: Graph M

Construction of a MAG

- Step-2:** For every bi-directed edge $A \leftrightarrow B$ in G ,
- (i) add $A \rightarrow B$ to M if A is an ancestor of B in G
 - (ii) add $A \leftarrow B$ to M if B is an ancestor of A in G
 - (iii) copy $A \leftrightarrow B$ to M if (i) and (ii) do not hold true

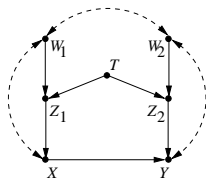


Figure: DAG G

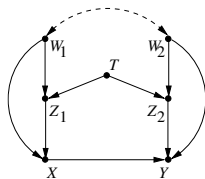


Figure: Graph M

Construction of a MAG

Step-3: For every pair of non-adjacent nodes A and B in G , connected by an *inducing path*,

- (i) add $A \rightarrow B$ to M if A is an ancestor of B in G
- (ii) add $A \leftarrow B$ to M if B is an ancestor of A in G
- (iii) add $A \leftrightarrow B$ to M if (i) and (ii) do not hold true

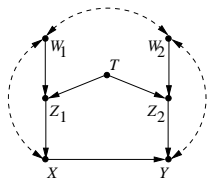


Figure: DAG G

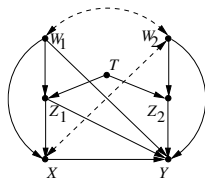


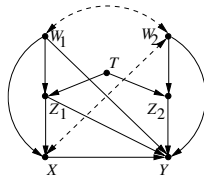
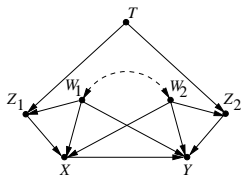
Figure: MAG M

Markov Equivalence

Theorem

Two graphs G_1 and G_2 are said to be Markov equivalent if their MAGs are Markov Equivalent

Are these MAGs Markov Equivalent?



Note: Markov Equivalence in MAGs are easier to check
 Complete criterion for determining Markov Equivalence of 2
 MAGs: [Ali, Richardson and Spirtes, 2009]

Reversing an edge in a MAG

Definition (Screened Edge)

[Tian,2005] An edge $X \rightarrow Y$ is a screened edge in a MAG if $Pa(Y) = Pa(X) \cup \{X\}$ and $Sp(Y) = Sp(X)$ ¹.

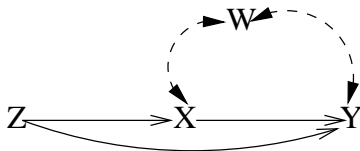


Figure: MAG with Screened Edge: $X \rightarrow Y$

¹Nodes X and Y are spouses, if they are connected by a bi-directed edge.

Reversing an edge in a MAG

Theorem (Tian,2005)

Let M be a MAG with edge $X \rightarrow Y$ and M' be a graph with edge $X \leftarrow Y$, otherwise identical to M . Then M' is a MAG that is Markov Equivalent to M if and only if $X \rightarrow Y$ is a screened edge in M .

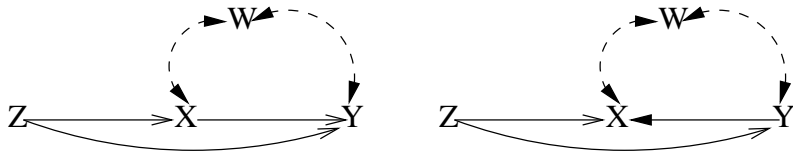


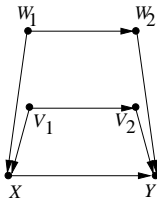
Figure: Markov Equivalent MAGs

Confounding Equivalence

Definition (Pearl and Paz, 2009)

Define two sets, T and Z as c -equivalent (relative to X and Y), written $T \approx Z$, if the following equality holds for every x and y :

$$\sum_t P(y|x, t)P(t) = \sum_z P(y|x, z)P(z) \quad \forall x, y$$



Note: C-equivalence is testable

Examples:

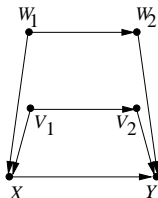
- ▶ $T = \{W_1, V_2\} \approx Z = \{W_2, V_1\}$
- ▶ $T = \{W_1, V_1\} \approx Z = \{W_2, V_2\}$
- ▶ $T = \{W_1, W_2\} \approx Z = \{W_1\}$
- ▶ $T = \{W_1, W_2\} \not\approx Z = \{W_2\}$

Necessary and Sufficient Condition for C-Equivalence

Theorem (Pearl and Paz, 2009)

Let Z and T be two sets of variables containing no descendant of X . A necessary and sufficient condition for Z and T to be c -equivalent is that at least one of the following conditions hold:

- ▶ $X \perp\!\!\!\perp (Z \cup T) \mid (Z \cap T)$ or
- ▶ Z and T are G -admissible²



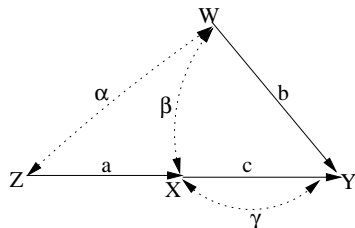
²satisfies back-door criterion

Examples:

- ▶ $T = \{W_1, V_2\} \approx Z = \{W_2, V_1\}$
- ▶ $T = \{W_1, V_1\} \approx Z = \{W_2, V_2\}$
- ▶ $T = \{W_1, W_2\} \approx Z = \{W_1\}$
- ▶ $T = \{W_1, W_2\} \not\approx Z = \{W_2\}$

Linear Models and Causal Diagrams

Assume all variables are normalized to have zero mean and unit variance.



$$Z = e_1$$

$$W = e_2$$

$$X = aZ + e_3$$

$$Y = bW + cX + e_4$$

$$\text{Cov}(e_1, e_2) = \alpha \neq 0$$

$$\text{Cov}(e_2, e_3) = \beta \neq 0$$

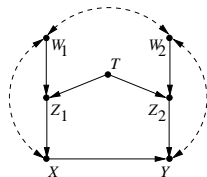
$$\text{Cov}(e_3, e_4) = \gamma \neq 0$$

Which parameters can be identified?

Vanishing Regression Coefficient

Definition

For any linear model for a causal diagram D that may include cycles and bi-directed arcs, the partial correlation $\rho_{XY.Z}$ must vanish if and only if node X is d-separated from node Y by the variables of Z in D [Spirtes et al., 1997b].



$$\blacktriangleright r_{TX.W_1Z_1} = 0$$

Find more

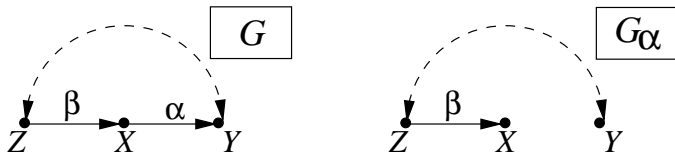
Single Door Criterion for Direct Effects

Theorem

Let G be any path diagram in which α is the path coefficient associated with link $X \rightarrow Y$ and let G_α denote the diagram that results when $X \rightarrow Y$ is deleted from G . The coefficient α is identifiable if there exists a set of variables Z such that :

- (i) Z contains no descendant of Y and
- (ii) Z d-separates X from Y in G_α

Moreover, if Z satisfies these two conditions, then α is equal to the regression coefficient $r_{YX.Z}$.



Instrumental Variables (IV)

Definition

A variable Z is an instrument relative to a cause X and an effect Y if:

- ▶ Z is independent of all error terms that have an influence on Y when X is held constant, and
- ▶ Z is **not** independent of X .

In linear systems, Causal effect of X on $Y = \frac{r_{ZY}}{r_{ZX}}$

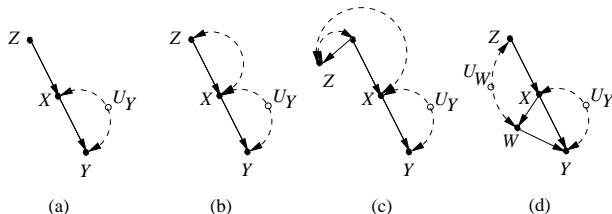


Figure: Z is an instrument in (a), (b) and (c) but not in (d)

Conditional Instrumental Variable

Definition (Brito & Pearl, 2002)

Z is an instrumental variable if \exists a set W such that:

- ▶ W contains only non-descendants of Y
- ▶ W d-separates Z from Y in the sub-graph G_α obtained by removing the edge $X \rightarrow Y$
- ▶ W does not d-separate Z from X in G_α

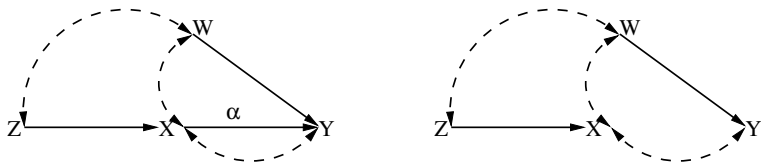


Figure: Graph G and corresponding subgraph G_α

Conditional Instrumental Variable

- ▶ Z is a conditional instrumental variable. Hence,
 $\alpha = \text{Causal effect of } X \text{ on } Y = \frac{r_{ZY.W}}{r_{ZX.W}}$
- ▶ W does not satisfy single-door criterion. So, α cannot be identified using single-door.

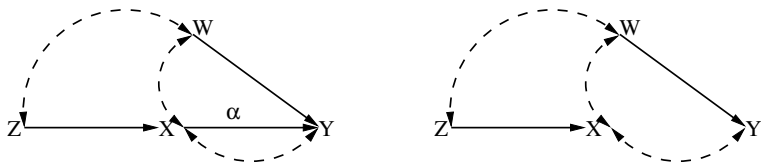
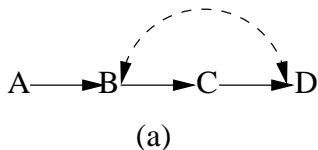


Figure: Graph G and corresponding subgraph G_α

Verma Constraints ([Tian and Pearl, 2002])



$$Q[\{B, D\}] = \sum_u P(b|a, u)P(d|c, u)P(u)$$

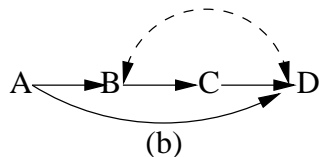
$$P_{v \setminus d}(d) = \sum_u P(d|c, u)P(u)$$

Also,

$$Q[\{B, D\}] = P(d|a, b, c)P(b|a)$$

$$P_{v \setminus d}(d) = \sum_b P(d|a, b, c)P(b|a)$$

$$\sum_b P(d|a, b, c)P(b|a) \text{ is independent of } a.$$



$$Q[\{B, D\}] = \sum_u P(b|a, u)P(d|a, c, u)P(u)$$

$$P_{v \setminus d}(d) = \sum_u P(d|a, u, c)P(u)$$

Also,

$$Q[\{B, D\}] = P(d|a, b, c)P(b|a)$$

$$P_{v \setminus d}(d) = \sum_b P(d|a, b, c)P(b|a)$$

$$\sum_b P(d|a, b, c)P(b|a) \text{ is **not** independent of } a.$$

Conclusions

Graphs are indispensable for:

- ▶ encoding causal assumptions
- ▶ identifying parameters and causal effects
- ▶ identifying testable implications

Go ahead and Exploit the Power of Graphs!

Thank You!